Reply to reviewer 1

This paper presents a significant enhancement of a Norwegian method for the estimation of extreme floods, based on an event-based rainfall-runoff simulation. It introduces a stochastic process for the assignment of the initial hydrological conditions before the simulated events, as well as for the intensity and the temporal dynamic of the simulated precipitation events. This method is compared to the initial method (which considers only a reference precipitation on given condition), and to a classical FFA. The presented method is interesting, both in terms of methodology and statistical results. It is well explored, with a detailed sensitivity analysis.

We thank the reviewer for the positive comment on the method and the very detailed review of the manuscript. However, the paper could be greatly improved by a better writing and more illustrations, particularly about the stochastic PQRUT which deserves a detailed step by step explanation of the simulation procedure (text and diagram), and also the probabilistic models for precipitation.

We agree that the explanation of the procedure can be improved. In the revised version, we have added step by step procedure:

1. Extract flood events for a given catchment and identify the critical storm duration
   For each season:
   2. Aggregate the precipitation data to match the critical duration for the catchment
   3. Extract POT precipitation events and fit a GP distribution
   4. Fit probability distributions for the initial discharge, soil moisture deficit and SWE values for the season
   5. Generate precipitation depth from the fitted GP distribution
   6. Disaggregate the precipitation depth to a 1 hour time step by matching the dates of the identified POT flood events (from step 3) to dataseries of precipitation with hourly timestep
   7. Sample a temperature sequence by matching the dates of the identified POT flood events (from step 1) to the dataseries of temperature with hourly timestep
   8. Sample initial conditions for snow water equivalent (SWE), soil moisture deficit and initial discharge from their distributions (step 5), accounting for co-variation using a multivariate normal distribution
   9. Simulate streamflow values using the calibrated PQRUT model for the sample event
10. Repeat steps 6.-9. 100 000 times
11. Estimate the annual exceedance probability from the total of 400 000 (100 000 for each season) samples using plotting positions

We have also included a diagram to illustrate the link between the different steps.

Regarding the sensitivity analysis, which is very important to understand the key factors and options of this new method, its writing also should be better organized and illustrated. It lacks a basic but important study of the impact of the random drawing (e.g. by performing 100 different simulations) and of the number of the simulated events on the extreme quantiles estimation. The later seems to be an issue here for high return periods.

These two issues are somewhat related. The effect of the random seed will be minimised if larger number of simulations are used. In addition, increasing the number of simulations will increase the robustness of the extreme quantiles. However, the number of simulations that are needed will depend on the return period of interest. Here we have considered a long return period, 1000-years, so this issue is indeed relevant. We have included two additional variables in the sensitivity analysis:
50 simulations using different random seeds
Length of simulation, of up to 400,000, 100,000 for each season (in increments of 40,000). It is not feasible to consider longer simulations due to computational times.

I would recommend a significant revision of this paper, mostly to improve its structure, its writing and the illustrations provided. Detailed comments/suggestions are provided to the authors in that follows.

We hope that our revisions will significantly improve the writing and the structure of the paper.

Abstract
For those not familiar with PQRUT, it could be added in the abstract that the stochastic PQRUT is an extension/evolution of the "standard" PQRUT routine, applied since many years in Norway (dates and references to be provided). The differences between the estimates can be up to 200\% for some catchments, which highlights the uncertainty in these methods. This is not a good message for hydrological engineering, a less pessimistic phrasing could be "[... 200\% for some catchments where the uncertainties of the compared methods are high and combine unfavourably]."

The reason, we did not include this information is that we wanted to emphasize that the method can be applied to any catchment with snowmelt/mixed flood regime. But as our study area is in Norway, we have revised the abstract as:

The estimation of extreme floods is associated with high uncertainty, in part due to the limited length of streamflow records. Traditionally, statistical flood frequency analysis and an event-based model (PQRUT) using a single design storm have been applied in Norway. We here propose a stochastic PQRUT model, as an extension of the standard application of the event-based PQRUT model, by considering different combinations of initial conditions, rainfall and snowmelt, from which a distribution of flood peaks can be constructed. . . . . The differences between the stochastic PQRUT and the statistical flood frequency analysis are within 50\% in most places. However, the differences between the stochastic PQRUT and the standard implementation of the PQRUT model are much higher, especially in catchments with a snowmelt flood regime.

§1 - Introduction Page 1, line 14: For example, floods with a 500-year return period are sometimes used to [...] As most of the estimates evoked in the paper are 100 or 1000-yr. floods, and example of the use of such quantiles in Norway could be useful.

We will revise the paragraph as follows:

The estimation of low-probability floods is required for the design of high-risk structures such as dams, bridges, levees, etc. For example, floods with a 100-year return period are sometimes required for the design of levees and the design and safety evaluation of high-risk dams requires the estimation of flood hydrographs for the 1000-year return period and, in some cases, floods with magnitudes of up to the Probable Maximum Flood (PMF). Page 1, line 17: Flood mapping also usually requires input hydrographs. This is also the case for dam safety assessment.

This sentence has been revised -See answer above

Page 1, line 21: When longer return periods are needed I guess the author means "longer than the record length", i.e. return period of 100 yr. or above.

Yes, this is correct. For clarity, this will be changed to:

When return periods that are longer than the observed record length are needed, the process requires extrapolation of the fitted statistical distribution. Page 2, line 6: have been shown to produce average errors between 27 and 70\% Please mention on what estimation this error is computed (observed quantiles or estimated ones, of which return period).

These errors are based on table 1 in the paper by Salinas et al 2013, which provides a comparative assessment of different studies in ungauged basins. The values are calculated using the RMSNE (root mean square normalised error) for Q100 (q100) and we have expressed them as percentage. We have revised as: As the physical processes in the catchments are usually not directly considered in the analysis, estimating the flood quantiles in ungauged basins using regression or geostatistical methods can produce average RMSNE (root mean square normalised error) values of between 27 and 70\% (Salinas et al., 2013), or even higher for the 100-year return period.
Page 2, line 32: they are computationally inefficient. . . Another writing could be "[. . . ] they are computationally demanding, as long continuous periods have to be simulated to estimate extreme quantiles".

Yes, this is in fact a better way to write this.

Page 3, line 6: millions of rainfall events can be sampled from the MEWP model. . . More exactly, millions of synthetic rainfall events can be generated, assigned to a probability estimated from the MEWP model, and inserted. . .

Yes, we agree with the revision.

Page 3, line 21: it requires the generation of a temperature sequence for the event I would add " a temperature sequence for the event, coherent with the simulated rainfall, and a snow water. . . "

We have revised this sentence as:

as it requires the generation of a temperature sequence for the event that is consistent with the rainfall sequence used and a snow water equivalent as an initial condition.

Page 3, line 22: The assumption of a fixed rate of snowmelt [. . . ] and a joint probability model needs to be considered Does it mean that a fixed snowmelt is usually added without consideration to the rest of the variables which will characterize the simulated event? What kind of joint probability model should be added?

To increase the clarity, these sentences has been rewritten as: The assumption of a fixed rate of snowmelt which is based on typical temperatures, as is often used in Norway for the single event-based design method, can introduce a bias in the estimates. The joint probability of both rainfall and snowmelt needs to be considered to obtain a probability neutral value (Nathan and Bowles, 1997).

Page 24, line 24: SEFM which has been applied in several USGS studies To my knowledge, I am not sure it is USGS (although SEFM is evoked in the USGS Bulletin 17C " Guidelines for Determining Flood Flow Frequency" of 2018), but several application of SEFM for dam safety studies have been delivered to USBR (US Bureau of Reclamation).

Yes, it should be USBR.

Page 3, line 34: due to the large uncertainty in both the event-based model and the statistical flood frequency analysis I am not comfortable with this writing. With two identical methods (say a classical FFA), but with two distributions fitted on two different samples, estimations would be different, and in that case this difference is completely linked to the uncertainties of the FFA (and mainly the sample uncertainty). But with different methods, these differences can also be produced by discrepancies between methods which should be treated per se, in order to assess the method themselves. So the interpretation of the difference should, in my opinion, not only rely on uncertainties.

This sentence has been deleted.

Page 3, line 35: To better understand the differences between these methods, a sensitivity analysis of the stochastic PQRUT is performed Here I have somehow the opposite comment from above: this sensitivity analysis of stochastic PQRUT is more about dealing with the uncertainties of stochastic PQRUT, not the differences between methods.

This is true, the sensitivity analysis will provide a better understanding of the uncertainty. However, the sensitivity analysis can also help us determine the reasons for the differences between the models. For example, the standard PQRUT assumes that fully saturated conditions are used. By testing the sensitivity of the model to the initial soil moisture deficit, we can check whether assuming fully saturated conditions contributes to the difference between these two methods.

We have revised this section as:

In order to understand the uncertainty and the differences with the standard PQRUT model, a sensitivity analysis of the stochastic PQRUT is performed by considering the effect of the initial conditions, model parameters and rainfall intensity on the flood frequency curve.
§2 - Stochastic event-based model Page 4, line 18: The study area consists of a set of 20 catchments A more logical phrasing could be "The study area in Norway, with a dataset of 20 catchments located throughout the whole country"

We have revised this sentence.

Page 5, line 6: for Krinsvatn, Lk and the area covered by marsh, M, is more than 10In the Table 1, Lk and M values are 9 and 1.1 %, respectively.

This is correct, we have revised as: but for Krinsvatn, the Lk is higher and the area covered by marsh, M, is 9

Page 5, line 28: the correlation of the method was found to be higher To which values the results of this disaggregation method have been correlated? Hourly or 3-hourly rainfall observations?

The disaggregated data were compared to the 3-hour observations in the study undertaken by Vormoor and Skaugen to produce the gridded disaggregated data. Page 5, line 29: simply dividing the seNorge data into eight equal parts To be clearer, I suggest "simply dividing the seNorge daily data into eight equal 3-hourly values".

The sentence has been revised as suggested. Page 5, line 29: disaggregated to a 1-hour time step using a uniform distribution to match the time resolution of the discharge data, although a 3-hour time step could also be used I don’t fully understand this. Was the 3-hour value affected randomly to one hour or divided into three? Why is it possible to use the 3-hour value with an hourly model?

The rainfall data was then divided into three equal parts, i.e. using a uniform distribution. The PQRUT model can be used with any timestep. Considering that the median catchment size is around 140 km2, a timestep of 3 hours should be sufficient for modelling the peak flows. As the model has previously been calibrated and regionalised using a 1-hour timestep, we decided to use this timestep (instead of 3-hours). A similar comment was raised by reviewer 2 and the section was revised to include these suggestions as well:

The HIRLAM atmospheric model for northern Europe has a 0.1 degree resolution (around 10 km2) and we used a temporal distribution of three hours. The HIRLAM data set was first downscaled to match the spatial resolution of the seNorge data and the precipitation of the HIRLAM data was rescaled to match the 24-hour seNorge data (Vormoor and Skaugen, 2013). Then, these rescaled values were used to disaggregate the seNorge data to a 3-hour time resolution. The method was validated against 3-hour observations, and the correlation of the method was found to be higher than that obtained by simply dividing the seNorge data into eight equal 3-hourly values (Vormoor and Skaugen, 2013). These datasets were further disaggregated to a 1-hour time step by dividing into three equal parts to match the time resolution of the streamflow data.

Page 6, line 1: remotely sensed data Assimilating remotely sensed data in a hydrological model is not an easy task, especially soil moisture. Any reference to provide that would apply to the context of this paper?

Using the remotely-sensed data currently available is probably not ideal as these data have a coarse spatial resolution (around 20km2). A review of the use of this data is given in Brocca et al (2017). This reference has been added to the revised version: Sources of these data can be e.g. remotely sensed data (see, for example, the review provided in Brocca et al. (2017)) or gridded hydrological models.

Page 6, line 2: the DDD hydrological model was used Please provide a reference which introduces this model.

The reference to the DDD model comes earlier in the manuscript, i.e. on p 4.

Page 6, line 7: exceeds 0.3 of its (dynamic) capacity Please define what is a dynamic capacity here.

In this case, dynamic refers to the concept that the volume of the saturated zone and unsaturated varies in time as explained in the previous sentence. This has been deleted.

Page 6, line 11: the so-called critical duration A reference can be provided here to define this concept: Meynink, W. J. Corderly, I. (1976). Critical duration of rainfall for flood estimation. Water Resources Research, 12(6), 1209-1214. Thanks for this, we have included this citation.
Page 6, lines 12-14: In order to determine [. . . ] had to be determined for each catchment I don’t find this sentence useful, considering what is written before and after it.

The sentence has been deleted.

Page 6, lines 14: flood events over a certain quantile threshold (0.9) were extracted. On what data this POT sample was extracted: daily or hourly discharges?

The sample is based on daily discharge, we have revised as: To determine the critical duration, flood events from the daily time series over a quantile threshold (in this case, 0.9) were extracted.

Page 6, lines 18-25: An alternative to this could be to study the correlation between the peak daily value and the precipitation of that day (which we could call P0), the sum P0 to P-1, P0 to P-2 and so on. . . When the correlation coefficient stops to increase significantly, it means that the correct length of the " precipitation window" is reached, thus the critical duration is estimated. This is likely to be more robust than studying the correlation between the peak daily discharge and the individual precipitations the days before.

This is an interesting suggestion. In this study, it is important to specify a critical duration in order to capture the “full” rainfall event producing the peak flow (otherwise the peak flow is likely to be underestimated). We implemented the procedure proposed by reviewer but we did not find much difference between the two methods. In fact, this alternative gives shorter critical durations in some cases (i.e. 1-day for Krinsvatn instead of 48 hours). As our study area consists of small catchments and the shortest window we are using is 1-day, the critical duration is then also 1 day (for 17 out of 20 catchments), rather than being longer.

Page 6, lines 22-24: In some catchments (mostly those having a snowmelt flood regime), no significant correlation was found between discharge and precipitation In that case, some processing of the flood is needed, e.g. only considering the " snowfree" seasons, or adding a threshold on the precipitation over the preceding days in the POT selection of floods. This could prevent using an arbitrary duration.

We have incorporated this and now only the relatively snow-free season (SON) have been considered.

Page 6, line 28: the sequence of the input data must be prescribed for the stochastic simulation What means " prescribed"? Is it generated? Is it randomly drawn from the observed sequences?

Yes, we have revised to “generated”.

Page 7, line 1: a Generalized Pareto distribution was fitted to the series of selected Events A figure with the corresponding fits and observations for the example catchments would be welcome.

We have added a return level plot that shows the fit to the observations for both GP and Exponential distributions.

Page 7, line 6: introduced in section ??

Section 2.1.2 , the reference has been corrected.

Page 7, line 7: Using the fitted Generalized Pareto (GP) distribution, precipitation depths were simulated Does it mean that probabilities where randomly drawn then the corresponding precipitation depths deduced from the fitted GPD? How many events are drawn?

This is correct, the sentence will be rewritten as:

The precipitation depths were generated (for 100,000 events) from the fitted GP distribution for each season. Originally, 160 000 (40 000 per season) simulations (around 10 000 to 20 000 years) were used, we have now increased this number to 400 000 events.

Page 7, line 8: a storm hyetograph was first sampled How is it sampled? I guess it comes from the hyetographs collection corresponding to the POT selection of precipitation events, but with what consideration to season, intensity, etc.?
Yes, it was sampled from the collection of hyetographs, the seasonality was considered but not the precipitation depth. This has been revised in the manuscript.

**Page 7, line 10: Pi and P are not defined in the disaggregation formula.**

We have revised as:

where $P_{\text{sim}}$ is the simulated 1-hour precipitation intensity, $P_d\text{sim}$ is the simulated daily intensity and $P_i$ is the 1-hour disaggregated SeNorge intensity

**Page 7, line 15: Output from DDD model runs. Have the DDD models been calibrated on local data? If yes, some words about the calibration method are welcome. I guess the DDD models are used at daily time-step, is it true?**

Yes, a daily time step was used. The following sentence has been added to the text: The model was calibrated for the selected catchments at a daily timestep using a MCMC routine (Petzoldt, 2010).

**Page 7, lines 21-25: The writing is not clear, and neither is the equation of the mixed distribution (what is $x$?). As far as I understood, it is about randomly switching between a trivariate (discharge, moisture, snow) and a bivariate (discharge, moisture) distribution depending on the probability of having snow on a given season. Is $p$ also drawn for the simulation?**

We agree that this was not clearly described, we have revised this section as follows: The probability $p$ for switching between the trivariate and bivariate distributions is based on the historical data for SWE higher than 0.

**Page 7, line 26: The correlation between the observed and simulated variables is shown in Figure 4. Apparently, $s_l$ is the soil moisture deficit. Contrary to SWE and $Q_{\text{obs}}$ which are "observable" variables, $s_l$ is linked to a model (here DDD). So it should be introduced, in relation with the DDD model structure.**

The soil moisture deficit is presented in Skaugen and Onof (2013). The soil moisture deficit is the difference between the volume of the unsaturated zone and the volume already present in the soil moisture zone.

**Page 8, line 5: for estimating design floods and safety check floods for dams in Norway. This type of application is perhaps documented in (Andersen, 1983), but this reference is not easily accessible on line, and is written in Norwegian, so a accessible reference documenting this type of application would be welcome.**

A reference to the NVE report is now provided:

The PQRUT model is a simple, event-based, 3-parameter model (Fig. 6) which is used, amongst other things, for estimating design floods and safety check floods for dams in Norway (Wilson et al., 2011).

**Page 8, line 14: The general procedures used for the PQRUT calibration are described in Filipova et al. Some details about this calibration would be welcome, e.g. which flood events sample is considered (is it the same as the one used in §2.2 for critical duration)?**

In the calibration, the 45 highest flood events were considered. This sample most likely overlaps with the events selected for the critical duration. The sentence has been revised as:

The PQRUT model was calibrated for the 45 highest flood events by using the DDS (Dynamically Dimensioned Search) optimization routine (Tolson and Shoemaker, 2007) and the Kling Gupta efficiency (KGE) criterion (Gupta et al., 2009) as the objective function.

**Page 8, lines 15-17: This additional parameter $l_p$ should be documented in the structure of the PQRUT model presented in the Figure 5. Furthermore, I am not sure that it can be considered as a parameter, more likely it is an internal state variable which vary from event to event.**

We agree in this case, $l_p$ can be considered as a state variable. The figure has been updated to include $l_p$.

**Page 8, line 18: the value of this parameter was set to the initial soil moisture deficit, estimated using DDD. This is an important assumption: it means that some internal variables of DDD (which ones, this is not documented) are used to estimate another one in PQRUT. This is far from obvious to accept for two very different models, running at different
time steps: what has be done to check this " compatibility"?

As we already discussed (see answer to Page 7, line 26), the DDD model is able to provide realistic values for soil moisture deficit. As we are interested in the antecedent soil moisture conditions and not the variation of the soil moisture deficit during rainfall event, the timestep is not of such importance. Other options would be to use soil moisture data from remote sensing or based on antecedent precipitation but these values are much less accurate. In addition, the soil moisture deficit values are not as important (as suggested by the sensitivity analysis) for high return periods.

Page 8, line 23: Cs is a coefficient accounting for the relation between temperature and snowmelt Properties It is usually called a " degree-day" coefficient (although used at a hourly time step here).

Both terms are used in literature but, as this is used in hourly time step, we prefer to use temperature index method

Page 8, line 30: The term under the bar should be " power to k" not be multiplied by k.

This seems to be correct- multiplied by k. The return period for the POT events is: T= 1/(k (1-P)) where k- is the number of events and P is the non-exceedance probability

Page 9, line3: These simulated events were compared with the POT flood events extracted from the observations At this point, I don’t clearly understand the simulation process. Some lines detailing the simulation process (sequence of random drawings, number of simulation, processing of events, etc.), as well as a diagram, are really necessary to the reader before entering into the analysis of the simulations.

The description has been revised (see previous answers). The simulated CDFs look affected by under-sampling above the 500 yr. return period (i.e. not enough simulations of this range), which interrogates the robustness of the 1000 yr. estimations which are assessed in the paper.

More simulations (400 000 instead of 160 000) have been used to address this issue (also see previous answers).

Page 9, line7: large variation in precipitation values Which duration is considered here? Daily? It is the total depth – in this case 24 or 48 hours. This is now also explained in the text. The comparison to the 100 yr. precipitation depths estimated thanks to the GP fit evoked in §2.3 would be useful.

A figure (fig 4) that shows the return level plots (which shows 100 -year return period) has been added to the revised version of the paper.

Page 9, line 14: even though fully saturated conditions are used in the event-based PQRUT model I don’t understand this: the lp variable (variable initial loss) has been introduced in §2.5 to depart from this fully saturated hypothesis.

Most commonly, fully saturated conditions are assumed for the standard PQRUT model. The reason for using the variable lp is to allow us to simulate flood events for which the initial conditions are not fully saturated.

Page 9, line 16: A sensitivity analysis was performed for the three test catchments Once again, the detailed protocol of this analysis deserves to be presented for a better understanding of the results. Some information is given in Table 3 but would deserve to be detailed in the text. A more logical " progression" of the different setups could be: 2, 3 (statistical hypothesis on precipitation), then 4 (temporal disaggregation), then 5,6,7 (simple hydrological assumptions) and finally 1 (PQRUT parameters). This would apply for the Table 3, as well as for the writing of §2.7

A similar comment was also raised by reviewer 2. In response to these comments we have included a table that illustrates the set up in the revised version of the manuscript.

Page 9, lines 24-28: I am not fully convinced by this explanation based on BFI. The sensitivity to initial loss should be linked to the possible values of initial loss in relation with the high quantiles of precipitation. I would be interested by looking at those values (maximum initial loss and 10, 100 and 1000 yr. precipitation) for the three catchments.

This is a good suggestion. We have included this analysis in the revised version: A reason for this is that for Øvrevatn, higher soil moisture conditions are associated with higher rainfall quantiles. For example, for Øvrevatn, precipitation depths with a
1000-year return period are associated with median soil moisture conditions of 37 mm, while for Krinsvatn, it is 30.8 mm and for Hørte, it is 16.7 mm.

Page 9, line 31: In addition, Krinsvatn shows high sensitivity to snowmelt. This is in contradiction with Page 9, line 10 (for Krinsvatn [. . . ] in most cases snowmelt does not contribute to the extreme floods). Any comment?

Krinsvatn shows a high sensitivity to excluding the snow component in the simulation. The reason is that the snowmelt is negative (there is snow accumulation). The sentence was revised as:

In addition, Krinsvatn shows a high sensitivity to the snowmelt component (21% higher) and also a step change in the frequency curve, even though the soil moisture deficit is higher. This can also be explained by the fact that the snowmelt contribution is negative (there is snow accumulation), as can also be seen in Table 2.

Page 10, line 7: Ovrevatn and Hørte showed sensitivity (28.9%) to the choice of the statistical distribution for modelling precipitation. A figure showing the precipitation distribution for each catchment (both observed and modelled by GP, and EXP) would be welcome to illustrate lines 7 to 10.

We have a add return level plot that shows the fit to the observations for both GP and EXP (see answer above)

§3- Comparison with standard methods Page 10, line 27: the standard implementation of the event-based PQRUT method. This is the first mention of such a "standard" implementation. I think this would deserve to be presented at the very beginning of the paper, which proposes a "stochastic PQRUT" being a significant enhancement from the "standard PQRUT". The context of this study would thus be better understood.

As of now, the Introduction provides a detailed overview of the methods for estimating extreme floods. Presenting the standard methods used in Norway in the introduction will narrow the scope, as potentially the international interest of this manuscript.

Page 10, line 29: the annual maximum series were extracted from the observed daily mean streamflow series. Why not using a GPD with the POT sample of floods extracted for the study of the critical duration?

The fitting of a GEV distribution to the AMAX series represents a standard implementation of the flood estimation guidelines in Norway (Midttømme et al. 2011). This is the reason why we used the AMAX series instead of the POT events.

Page 10, line 31: to obtain instantaneous peak values, the return values were multiplied by empirical ratios, obtained from regression equations. Here I don’t understand why the POT flood events extracted from observations (shown in the plots of the Figure 6) has not been used to fit either a GPD, or a GEV after extraction of annual maxima. More comments about this would be welcome.

Much longer series of data are available at daily timestep than at sub-daily timesteps, as technology making sub-daily series widely available was only introduced during the 1980s, whereas many daily records are over 100 years in length. Fitting a GPD distribution to the instantaneous peak flows and using this model to predict the 100-year return period will involve much higher uncertainty.

Page 11, line 11: obtained from growth curves based on the 5-year return period value. If I understand properly, the shape of the design hyetograph is based on the growth curves considered at the 5 yr. return period. Are the ratios between the different duration values at this return period deduced from empirical distribution, or inferred from a fitted distribution? Later on, this must be scaled to define a 100 or 1000 yr. hyetograph. What precipitation distribution (duration and model) are these extreme values deduced from?

The Gumbel distribution is used to derive the growth curves, while the ratios between the different durations are derived from an empirical distribution following a procedure developed by NERC in the UK in the 1970s and later applied in Norway, based on Norwegian data. This section has been revised: The standard implementation of PQRUT involves using a precipitation sequence that combines different intensities, obtained from growth curves based on the 5-year return period value fitted using a Gumbel distribution while the ratios between the different durations are derived from empirical distribution (Førland, 1992).
Page 11, line 15: The performance of the three models was validated by using two different tests. In that case, dealing with 100 or 1000 yr. flood estimations, it’s more about "comparing different approaches".

Even though the uncertainty is high for these return periods, a check that the data is within the confidence interval can be used as a validation (e.g. Lamb et al. 2016).


Page 11, line 20: As discussed, due to the difficulty in assigning initial conditions for the event-based PQRUT model I don’t understand this sentence, and to which discussion it refers.

This refers to the fact that fully saturated conditions are used in the standard implementation of the PQRUT model. The following sentence was added: The standard implementation of the event-based PQRUT model was not evaluated based on QS as initial conditions could not be assigned for low return periods. As this model is usually used to calculate high quantiles (Q100 or higher), fully saturated conditions are assumed for its implementation.

Page 11, line 22: The regional equations were used. Which regional equations? For PQRUT parameters?

Yes, we used the regional equations for the PQRUT parameters. This is now correct in the revision.

Page 11, line 25: equation of QS + observed probabilities (Qobsi) are calculated using Gringorten positions for the POT series. The POT series are used here, contrary to the daily (transposed to peak) annual maximum values that have been fitted in the statistical approach. Another option (already mentioned in my remark for page 10, line 29) could be to fit the statistical method on the POT sample, which would have allowed to keep it as a "benchmark" method, given more sense to the comparison presented (or conversely using the "peak-from-daily" observations for the QS calculation).

This is a good point; the daily flows were used for the QS calculation. We have revised the description of the method: In Eq. 5 the observed probabilities (Qobsi) are calculated using Gringorten positions for the peak AMAX series that were derived from the daily values. The modelled probabilities that correspond to the observed events are calculated by using the statistical flood frequency analysis and the Stochastic PQRUT model, as described previously.

Page 11, line 30: The results vary between catchments as shown in fig 8. I don’t find this figure very useful, the reader is unable to interpret the coloured dots. An alternative, aside the boxplots, could be some scatter plots (statistical Q100 and Q1000 v/s standard and stochastic PQRUT, statistical QS v/s standard and stochastic PQRUT, etc.).

We have included a figure that shows the QS for each catchment.

Page 11, line 32: we can conclude that the performance of the standard PQRUT model is poorer than the performance of the statistical flood frequency analysis and the stochastic PQRUT model. The results which ground this conclusion are not explicitly presented. The only clue given to the reader is the Figure 8 which only presents the distribution of QS scores for FFA and stochastic PQRUT. The results, in terms of QC score as well as confidence interval, should be presented in a table and in an adequate figure.

In addition to the figure that shows the QS, we have included a figure that shows the number of models that are within the confidence interval for Q100 for each catchment.

Page 12, line 3: The violin plots (fig. 9). See remarks on Figure 9 below.

Figure 9 shows both violin plots and boxplots (overlayed in gray). In order to increase the readability of the figure, only the boxplots are now plotted.

Page 12, line 7: Reasons for this may be that higher precipitation intensity or snowmelt is used. To assess this, the values of the reference hyetographs used in standard PQRUT deserve to be presented and compared to the simulated precipitation values of stochastic PQRUT (like the values of the Table 2 for Q100).
The results for Q1000 obtained from the stochastic PQRUT are now lower (after increasing the number of simulations).

Page 12, line 8: the absolute differences between the two methods are larger in catchments with lower temperature (fig. 9) I wonder how this can be deduced of illustrated by Figure 9, it is more likely somehow in Figure 10.

Apologies for this error, the figure numbers have now been updated.

Page 12, line 17: which might be due to the uncertainty in estimating the parameters for the GEV distribution I don’t understand this interpretation which appears rather quick and subjective to me. We agree, this section has been deleted in the revised manuscript.

Page 12, lines 18-21: This using of the study of Rogger et al. (2012) is off topic for me here, as it is based on Gumbel, whereas the FFA is done here with GEV, which is more flexible.

This is true, this is the reason that in the paper we specifically discuss the fact that the Gumbel distribution was used in the study of Rogger et al. (2012). This section has been deleted.

§4- Conclusions Page 13, line 10-15: Another modelling option could be to run the event-based simulation with the DDD model, already used for the initial condition. In that case, an hourly version of DDD should also be calibrated (with local observations or regionally), in compatibility with the daily version used for initial conditions. I am not fully aware of the potential difficulties of this, but it would be a more homogeneous approach in terms of hydrological modelling. Any comment about this?

Yes, this is a possibility. However, in large catchments it is not as important to use a subdaily timestep, as the peak and daily flows are similar.

Page 13, line 28: easily incorporate the uncertainty associated with this choice This is a very good remark: the stochastic process here adequately models a variable which, when represented in a deterministic way (i.e. fixed initial conditions), appears as highly uncertain.

Yes, this has also been discussed in several other studies.

Page 13, line 31: based on an assessment of the uncertainty characterizing the individual methods This is an interesting suggestion, but it has to be added that a proper expression of uncertainty for a rather sophisticated method like stochastic PQRUT is far from trivial, and is still to be investigated. . .

This sentence has been deleted.

Tables Table 1 Units are missing, as well as legend of the columns in the caption. Table 2: Units are missing, precipitations could be rounded to the next mm. For Krinsvatn, the probability to find the Q100 events in one season or the other could be provided.

The two tables have now been revised.

Figures Figure 3: Not very informative, re-scaling storm hyetograph is not something difficult to understand. A set of different "typical" hyetographs could instead be presented for the three catchments, ideally illustrating the potential diversity of storm dynamic. A new figure was added. Figure 4: "for Krinsvatn catchment" could be added in the caption, as well as the number of observed and simulated events.

The figure has been updated.

Figure 6: The remarkable return periods (10, 100 and 1000 yr.) should be distinguished in the plots (by a bolder vertical line for example).

The figure has been updated.

Figure 7: There are too many distributions in the plots, their interpretation is not easy. Two plots could be edited for each catchment, having for example only the "calibrated" simulation in common. An uncertainty band around the "

10
calibrated" simulation would be useful to assess the intrinsic uncertainty of the simulation process.

The figure has been updated.

Figure 8: See comment of page 11, line 32.

This will be added.

Figure 9: I am not convinced by the usefulness of the violin plots here considering the limited number of values per scores (20 catchments). Box plots with outliers would have been sufficient and more readable. Captions of the methods sometimes overlap.

Reply to reviewer 2

General Comments: As I understood, the main purpose of the work is to propose a methodology to overcome the limitations of more commonly applied event based modelling for flood frequency estimations by a stochastic modelling of preconditions, including SWE, and meteorological input. The individual modelling of the different aspects are described in the manuscript, however, it is hard to follow how the different parts are connected. A preceding sub-section with a less detailed step by step explanation of methodology, maybe including a schematic illustration (inputs/ models/ methods / output), could help to better explain the methodology.

A similar comment has been made by reviewer 1. We have added a step by step description and also included a diagram to illustrate the link between the different steps.

For the validation of the disaggregation procedure the disaggregated data were compared against hourly station data. Is this correct? I would be interesting to see how well the disaggregation procedure was performing (For example showing a obs-sim, QQ-plot). It is stated that it works better than equal divisions which is not surprising. What is the advantage of the further equal division to 1h if it is stated that 3-hours are already enough? Further, it is not obvious why the gridded seNorge.no Data are matched to the HIRLAM data if they are in the needed temporal resolution already?

The HIRLAM data is a hindcast dataset with a spatial resolution of around 10 km2 and a temporal resolution of 3 hours. The gridded seNorge data is obtained by triangulation of the observed rainfall dataseries; it has a spatial resolution of 1km2, and a temporal resolution of 24-hours. As the HIRLAM data has a higher temporal resolution than the seNorge data, the HIRLAM data was used to disaggregate the seNorge data to a 3-hour timestep. The performance, including the validation of the disaggregation procedure, is described in Vormoor and Skaugen (2013).

For the work presented here, the precipitation data were further disaggregated to a 1-hour time step by dividing into three equal parts. This was simply done for convenience, as the PQRUT model has previously been calibrated relative to 1-hour streamflow data and similar climate input data (i.e. 1-hour data derived from 3-hour data by dividing into three equal parts). A similar comment was raised by reviewer 1, and this section will be revised to include the above clarifications.

We have revised the paragraph as:

The HIRLAM data is a hindcast dataset with a spatial resolution of around 10 km2 and a temporal resolution of 3 hours. The HIRLAM data was first downscaled to match the spatial resolution of the seNorge data and the precipitation of the HIRLAM data was rescaled to match the 24-hour seNorge data (Vormoor and Skaugen, 2013). Then, these rescaled values were used to disaggregate the seNorge data to a 3-hour time resolution. The method was validated against 3-hour observations, and the correlation of the method was found to be higher than that obtained by simply dividing the seNorge data into eight equal 3-hourly values (Vormoor and Skaugen, 2013). These datasets were further disaggregated to a 1-hour time step by dividing into three equal parts to match the time resolution of the streamflow data.

A 1000 years event is extrapolated from daily observation series (length not further specified). Furthermore the results are then multiplied by empirical factors, to match sub-daily peak flows. I am not aware of the engineering practice in Norway, however, I am not sure about the meaning of the results by this extreme extrapolation and at least this
should be critically discussed.

The fitting of an extreme value distribution to estimate the return level for periods longer than the length of a time series is a standard procedure, both in hydrological investigations and in engineering practise. As suggested by the reviewer, the uncertainty of the estimates does increase significantly for longer return periods, relative to the length of record. The length of the daily streamflow series considered here, however, justifies the use of an ‘at-site’ (cf. a regional) flood frequency analysis as the minimum length is 31 years, while the median is 65 years of data. The following sentences will be added to the text:

In addition, the length of the daily streamflow series justifies the use of at-site flood frequency analysis (Kobierska et al., 2017); the minimum length is 25 years, while the median is 65 years of data. However, it is expected that the uncertainty will be high when the fitted GEV distribution is extrapolated to a 1000-year return period. The 1000-year return period is used here, however, as it is required for dam safety analyses in Norway (e.g. Midttømme, et al., 2011; Table 1). More robust, but potentially less reliable, estimates could be obtained using a 2-parameter Gumbel, rather than a 3-parameter GEV distribution (Kobierska et al., 2017).

The sensitivity analysis is interesting, however, also confusing including Figure 7 and Table 3. It is not obvious on what basis the percentage difference is calculated. This is also not clear in the follow up comparison of the methods. What exactly is the calibrated model? Also the section misses an explanation of the shaded area which is prominently displayed in Figure 7. Furthermore, the different precipitations settings tested are not well explained. A table, summarizing the different tested aspects, would help to guide the reader.

Thank you for these comments and suggestions. The shaded area represents the simulations based on the 5% and 95% confidence intervals for the regression equations for PQRUT. We have included a table to describe the set up and we have also revised the paragraph that describes the sensitivity analysis as follows:

A sensitivity analysis was performed for the three test catchments, Hørte, Øvrevatn and Krinsvatn, in order to determine the relative importance of the initial conditions, precipitation, the parameters of PQRUT, the effect of the random seed and length of simulation on the flood frequency curve. To test the sensitivity of the model, we have used several different model runs and calculated the percentage difference of each of these model runs relative to the standard model setup, as shown in Fig.8. More detailed information on the set up is given in Table 3. As these catchments are located in different regions and exhibit different climatic and geomorphic characteristics, we hypothesize that the flood frequency curve will be sensitive to different parameters and hydrological states, as well as local climate and catchment characteristics. The results are summarised in Table 4.

The Figures and especially the captions should be improved, as they are often not self-explanatory. This includes also missing units, labels and abbreviations. Maybe consider a professional language proof reading.

In the revised version of the article, we have tried to improve the language usage.

Specific Comments: The abbreviation PQRUT, used from beginning (abstract), is not introduced on page 4 or rather page 8. Please declare the meaning of PQRUT first time mentioned.

The abbreviation PQRUT comes from P-precipitation, Q- discharge and RUT -routing, this has been updated.

The characterizations of catchments and chosen abbreviations are introduced on P4 and repeated later (P5, l5) without brackets (e.g. “sparse vegetation over tree line (B)” and “sparse vegetation over tree line B”). Either use brackets throughout the manuscript or only use the abbreviation. Additionally by choosing more selfexplanatory abbreviations or using full words (eg. forest; marsh), would be easier to understand, especially in Table 1.

The abbreviation are now explained in the table caption.

P.5 l.15: The last sentence does not contain important informations and could be omit

We prefer to keep this sentence as it gives useful information on how the data is derived and increases the reproducibility of the study.
P.6 l.1: The addition “,which can be used for modelling in ungagged basins.” could be omitted, as it seems not connected to the procedure.

This sentence has been deleted.

P.6 l.2: A citation should be added to the DDD model or the corresponding R-package.

The reference is provided earlier in the manuscript, p4.

P.6 l.11: In my opinion, the meaning of the “critical duration” rather reflects the link between the duration and intensity of precipitation “events” of a certain probability, than to ensure the modelling of the complete flood hydrograph.

This is a good point, the sentence will be revised to:

When simulating flood response with an event-based model, it is important to specify the so-called critical duration (Meynink, W. J., Cordery, I. 1976) to ensure that the flood peak is correctly modelled. The critical duration is an important factor which effectively links the duration and the intensity of precipitation events of a given probability.

P.6 l.32: “individual risk seasons could have been defined”. One wonders why it was not done? If not so important for the result, please consider to omit this half sentence.

This sentence has been deleted. We used this season definition to match the seasonal definition used by the Norwegian Meteorological Institute.

P.8 l.12-17: Please check grammar and style of the section.

This section has been revised to also address the issues raised by reviewer 1 as follows.

The PQRUT model was calibrated for the 45 highest flood events by using the DDS (Dynamically Dimensioned Search) optimization routine (Tolson and Shoemaker, 2007) and the Kling Gupta efficiency (KGE) criterion (Gupta et al., 2009) as the objective function. An additional variable, the soil deficit, \( \ell_p \), was introduced to account for initial losses to the soil zone. The reason for this is that, even though fully saturated conditions are assumed when the model is used to estimate PMF or other extreme floods with low probabilities, the model needs to account for initial losses when actual (more frequent) events are simulated. This procedure is described in more detail in Filipova et al. (2016). In addition, regional values can be used in ungauged or poorly catchments (Andersen et al., 1983; Filipova et al., 2016).

P.8 l.28: Was there a specific reason for using the “Gringorten plotting” position?

The Gringorten plotting positions provide unbiased quantile estimates for the Gumbel distribution. In this case, we don’t know the distribution. However, the difference between the plotting positions is usually higher for the low and high quantiles. As reviewer 1 suggests we have increased the number of simulations. This means that differences derived from plotting position formulas will be relatively small when estimating the 1000-year return period.

P.12 l.3: A more detailed explanation what exactly is analyzed here is missing.

The sentence will be revised to: A comparison of the stochastic PQRUT with the standard methods for flood estimation shows that there is a large difference between the results of the three methods for both Q100 and Q1000 (Fig. 12 and 13).

P.12 l.25: Maybe the catchment steepness should be introduced in the section “study area”.

We have now added the catchment steepness to the “study area” section.

P.13 l.25: Why is it peak to volume? I thought it is daily mean to daily max discharge?

This refers to converting the daily volume (obtained from the daily mean) to the peak value.

Table 1: Missing units. Furthermore, the variables could be sorted and clustered more logically (e.g. temperature and precipitation; Q and AMAX).

Thanks, the units have now been included.

Table 3: Why are 100 values sampled? Does T mean the threshold Parameter \( \text{T} \)?
For the sensitivity analysis we used 50 samples, as larger number will increase the computational time. We assume that this number is sufficient to calculate the intervals. Trt refers to the parameter of the PQRUT model, thanks for spotting this error.

**Figure 2: Labels and units are missing**

This has been corrected for the revised version of the manuscript.

**Figure 3: Labels and units are missing**

This has been corrected for the revised version of the manuscript.

**Figure 7: is confusing because of the large number of different colored lines. Maybe two plots can help to distinguish between the different aspects as for example the precipitation input and other aspects. The legend is confusing as well. GDP was fitted to what? Y-Axes should start at 0, x-axes missing a label and to be consistent with the rest of the work it should not exceed 1000.**

This issue was also raised by reviewer 1 and the figure will be improved and revised based on the newer, longer simulations.

**Figure 8: It is impossible to distinguish between 20 colors. Do the colors have any meaning? If they should be recognizable, numbering would be a better option. The numbers could then also be used in Figure 10, so the link between the performance of the model and the results are given.**

The colors just represent different catchments but also as reviewer 1 suggests. We have replaced this figure with a boxplot that shows the performance of the methods at each catchment.

**Figure 9: What exactly is shown in the plots. Please add a more detailed explanation.**

The figure has now been replaced. Instead, now we are using boxplots to show the differences in the performance of the three methods.

**Figure 10: The scale “percentage difference” should be unambiguous. The base of the “difference” should be clarified.**

We have revised the description of the figure as:

Results of the comparison between stochastic PQRUT, PQRUT and GEV for the values of the 1000-year return level. The absolute differences (calculated by dividing the estimate by the average of all models) are correlated with catchment properties. Positive correlations are given by red and negative by blue.
A stochastic event-based approach for flood estimation in catchments with mixed rainfall/snowmelt flood regimes

Valeriya Filipova\(^1\), Deborah Lawrence\(^2\), and Thomas Skaugen\(^2\)

\(^1\)University of Southeast Norway, INHM, Gullbringvegen 36, 3800 Bø, Norway, e-mail: valeriya.filipova@usn.no

\(^2\)Norwegian Water Resources and Energy Directorate, P.O. Box 5091 Maj., N-0301 Oslo, Norway, e-mail: dela@nve.no

**Correspondence:** Valeriya Filipova (valeriya.filipova@usn.no)

**Abstract.** The estimation of extreme floods is associated with high uncertainty, in part due to the limited length of streamflow records. Traditionally, statistical flood frequency analysis or and an event-based model (PQRUT) using a single design storm have been applied. We propose here an alternative, stochastic in Norway. We here propose a stochastic PQRUT model, as an extension of the standard application of the event-based modelling approach. The stochastic PQRUT method involves Monte Carlo procedure to simulate PQRUT model, by considering different combinations of initial conditions, rainfall and snowmelt, from which a distribution of flood peaks can be constructed. The stochastic PQRUT was applied for 20 small and medium-sized catchments in Norway and the results show good fit to the observations give good fits to observed peak-over-threshold series. A sensitivity analysis of the method indicates that the soil saturation level is less important than the rainfall input and the parameters of the PQRUT model for flood peaks with return periods higher than 100 years, and that excluding the snow routine can change the seasonality of the flood peaks. Estimates for the 100- and 1000-year return level based on the stochastic PQRUT model are compared with results for a) statistical frequency analysis, and b) a standard implementation of the event-based PQRUT method. The differences between the estimates can be up to 200% for some catchments, which highlights the uncertainty in these methods. Stochastic PQRUT and the statistical flood frequency analysis are within 50% in most places. However, the differences between the stochastic PQRUT and the standard implementation of the PQRUT model are much higher, especially in catchments with a snowmelt flood regime.

1 **Introduction**

The estimation of low-probability floods is required for the design of high-risk structures such as dams, bridges, levees, etc. For example, floods with a 500-year 100-year return period are sometimes used to evaluate the risk of scour to bridges (Ries, 2007) required for the design of levees and the design and safety evaluation of high-risk dams requires that the estimation of flood hydrographs for the 1000-year return period and, in some cases, floods with magnitudes of up to the Probable Maximum Flood (PMF) are estimated. An overview of design flood standards for reservoir engineering in different countries is provided in Ren et al. (2017). Flood mapping also usually requires input hydrographs for flood events with return periods of up to 1000 years. Methods for estimating these floods can be generally classified into three groups: 1) statistical flood frequency analysis; 2) the single design event simulation approach; and 3) derived flood frequency simulation methods.
At gauged sites, statistical flood frequency analysis involves fitting a distribution function to the annual maxima or peak over threshold flood events and calculating the quantile of interest. When longer return periods are needed, the process requires extrapolation of the fitted statistical distribution, which introduces a high degree of uncertainty due to the number of limited observations relative to the estimated quantile (e.g. Katz et al., 2002). Significant progress has been made in methods for reducing this uncertainty by incorporating historic or paleo-flood data (Parkes and Demeritt, 2016), where available. Another way to “extend” the hydrological record in order to reduce the uncertainty is to combine data series from several different gauges by identifying pooling groups or hydrologically similar regions, where this is possible. It has been found, however, that the identification of such hydrological regions can be difficult in practice (Nyeko-Ogiramoi et al., 2012). The application of statistical flood frequency analysis in ungauged basins is also problematic. As the physical processes in the catchments are usually not directly considered in the analysis, estimating the flood quantiles in ungauged basins using regression or geostatistical methods have been shown to produce average errors can produce average RMSNE (root mean square normalised error) values of between 27 and 70% (Salinas et al., 2013), or even higher for the 100-year return period. In addition, the complete hydrograph is often needed in practice. Although multivariate analysis of flood events (e.g. flood peaks, volumes and durations) can be used to generate hydrographs for specific return periods, the methods are not easily applied (Gräler et al., 2013).

The second method for extreme flood estimation is the design event approach in which single realizations of initial conditions and precipitation are used as input in an event-based hydrological model. Another feature of the approach is that when event-based models are used, a critical duration defined as the duration of the storm that results in the highest peak flow needs to be set. Advantages included. Two advantages of this method over statistical flood frequency analysis is that rainfall records are often widely available (e.g in the form of gridded datasets) and that the event hydrograph is generated in addition to the peak magnitude. This approach has been traditionally used due to its simplicity (e.g. Kjeldsen, 2007; Wilson et al., 2011). However, its application often involves the assumption that the simulated flood event has the same return period as the rainfall used as input in the hydrological model. This assumption is not realistic and, depending on the initial conditions, the return period of the rainfall and the corresponding runoff can differ by orders of magnitude (e.g. Salazar et al., 2017). A reason for this is that flood events are often caused by a combination of a high level of saturation, rainfall and snowmelt, and factors, such as a high degree of soil saturation in the catchment, heavy rainfall and seasonal snowmelt. A joint probability distribution, therefore, needs to be considered if one is to fully describe the relationship between the return period of rainfall and of runoff.

The third possible approach is the derived flood frequency method in which the distribution function of peak flows is derived from the distribution of other random variables such as rainfall depth and duration and different soil moisture states. Although a statistical distribution of flow values or their plotting positions are then used to calculate the required quantiles, as in conventional flood frequency analysis, a hydrological model can be used to derive discharge values and thus extend and complement simulate an unlimited number of discharge values under differing conditions, thus extending and enhancing the observed discharge record. The Under very stringent assumptions, the derived distribution can actually be solved analytically by using a simple rainfall-runoff model (e.g. a unit hydrograph), assuming independence between rainfall intensity and duration,
and considering only a few initial soil moisture states. However, because of these simplifying assumptions, the method can produce poor results (Loukas, 2002). For this reason, methods based on simulation techniques are most often used, and these range from continuous simulations to event-based simulations with Monte Carlo methods.

In the continuous simulation approach, a stochastic weather generator is used to simulate long synthetic series of rainfall and temperature, which serve as input in a continuous rainfall-runoff model. The resulting long series of simulated discharge are then used to estimate the required return periods, usually using plotting positions (e.g. Calver and Lamb, 1995; Camici et al., 2011; Haberlandt and Radtke, 2014). A disadvantage of these methods is that they are computationally inefficient in that long periods between extreme events are also simulated demanding, as long continuous periods need to be simulated to estimate the extreme quantiles. Several newer methods, therefore, use a continuous weather generator coupled with an event-based hydrological model. For example, the hybrid-CE (causative event) method uses a continuous rainfall-runoff simulation to determine the inputs to an event-based model (Li et al., 2014). Another disadvantage, however, of continuous simulation models is that stochastic weather generators require the estimation of a large number of parameters (e.g. Onof et al., 2000; Beven, Keith & Hall, 2014). In addition, models such as the modified Barlett-Lewis rectangular pulse model have limited capacity to simulate extreme rainfall depths, which can lead to an underestimation of runoff (Kim et al., 2017). In order to avoid the limitations of the continuous weather generators, the semi-continuous method SCHADEX (Paquet et al., 2013) uses a probabilistic model for centered rainfall events (MEWP; Garavaglia et al. (2010)), identified as over threshold values that are larger than the adjacent rainfall values. Using this approach, millions of synthetic rainfall events can be generated, assigned a probability estimated from the MEWP model, and inserted directly into the historic precipitation series to replace observed rainfall events. In this manner, the SCHADEX method is similar to the hybrid-CE methods because a continuous hydrological model long-term hydrological simulation is used to characterize observed hydrological conditions and synthetic events are only inserted into the precipitation record for periods selected from the observed record. Despite the many advantages of the hybrid-CE and SCHADEX methods over continuous simulation methods, they still require sufficient data for the calibration of the hydrological model, modelling of the extreme precipitation distribution and for ensuring that an exhaustive range of initial hydrological conditions are sampled during the simulations.

Another method for derived flood frequency analysis is the joint probability approach (e.g. Muzik, 1993; Loukas, 2002; Svensson et al., 2013; Rahman et al., 2002). In this approach, Monte Carlo simulation is used to generate a large set of initial conditions and meteorological variables, which serve as input to an event-based hydrological model. This approach requires that the important variables are first identified and any correlations between the variables are quantified. Most often, the random variables that are considered are related to properties of the rainfall (intensity, duration, frequency) and to the soil moisture deficit. Some of these methods, such as the Stochastic Event Flood Model (SEFM, (Schaefer and Barker, 2002)), similarly to SCHADEX, require the use of a simulation based on a historical period to generate data series of state variables from which the random variables are sampled. Although the contribution of snowmelt can be important in some areas, it is rarely incorporated as it requires the generation of a temperature sequence for the event that is consistent with the rainfall sequence used and a snow water equivalent as an initial condition. The assumption of a fixed rate of snowmelt which is based on typical temperatures, as is often used in Norway for the single event-based design method, can introduce a bias in the
estimates and a joint probability model. The joint probability of both rainfall and snowmelt needs to be considered to obtain a probability neutral value (Nathan and Bowles, 1997). One of the few methods that incorporates snowmelt is the SEFM which has been applied in several USGS USBR (US Bureau of Reclamation) studies and uses the semi-distributed HEC-1 hydrological model (Schaefer and Barker, 2002). Considering that, most often, simple event-based hydrological models are used (e.g. unit hydrograph), the joint probability approach is particularly advantageous in ungauged catchments or data-poor catchments, where the use of parsimonious models is preferred.

Even though methods for derived flood frequency analysis are becoming more commonly used in practice as they can provide better estimates of the high flood quantiles (e.g. Australian Rainfall and Runoff 2016 (?), SCHADEX method), this method has not yet been established in Norway. The purpose of this study is hence to develop a derived flood frequency method using a stochastic event-based approach to estimate design floods, including those with a significant contribution from snowmelt. In this way, the results for any return period can be derived, taking into account the probability of a range of possible initial conditions. A sensitivity analysis is then performed to understand the uncertainty in the stochastic PQRUT model and establish the relative roles of several factors, including rainfall model, snowmelt, initial soil moisture parameters of the model and the length of the simulation. The results are then compared with results from an event-based modelling method based on a single design precipitation sequence and assumed initial conditions and with statistical flood frequency analysis of the observed annual maximum series for a set of catchments in Norway. The methods give different results in many of the catchments due to the large uncertainty in both the event-based model and the statistical flood frequency analysis. To better understand the differences between these methods, a sensitivity analysis of the stochastic PQRUT is performed by considering the effect of the initial conditions, model parameters and rainfall intensity on the flood frequency curve for the 100- and 1000-year return period.

2 Stochastic event-based flood model

The stochastic event-based model proposed here involves the generation of several hydrometeorological variables: precipitation depth and sequence, temperature, the temperature sequence during the precipitation event, antecedent discharge, the initial discharge, and the antecedent soil moisture conditions and antecedent snow water equivalent. A simple 3-parameter flood model PQRUT (Andersen et al., 1983) is used to simulate the streamflow hydrograph for a set of randomly generated selected conditions based on the hydrometeorological variables these hydrometeorological variables. After this procedure is completed 100,000 times for each of the four seasons, the results are combined and a flood frequency curve is constructed from all of the simulations using their plotting positions. As the method requires initial values for soil moisture and snow water equivalent, i.e variables which generally cannot be sampled directly from climatological data and which depend on the sequence of precipitation and temperature over longer periods, the Distance Distribution Dynamics (DDD) hydrological model (Skaugen and Onof, 2014) was calibrated and run for a historical period to produce a distribution of possible values for testing the approach. The method (also shown in Fig. 1) can be outlined in summary form as follows:
1. Extract flood events for a given catchment and identify the critical storm duration. For each season:

2. Aggregate the precipitation data to match the critical duration for the catchment

3. Extract POT precipitation events and fit a GP distribution

4. Fit probability distributions for the initial discharge, soil moisture deficit and SWE values for the season

5. Generate precipitation depth from the fitted GP distribution

6. Disaggregate the precipitation depth to a 1 hour time step by matching the dates of the identified POT flood events (from step 3) to dataseries of precipitation with hourly timestep

7. Sample a temperature sequence by matching the dates of the identified POT flood events (from step 1) to the dataseries of temperature with hourly timestep

8. Sample initial conditions for snow water equivalent (SWE), soil moisture deficit and initial discharge from their distributions (step 5), accounting for co-variation using a multivariate normal distribution

9. Simulate streamflow values using the calibrated PQRUT model for the sample event

10. Repeat steps 6.-9. 100 000 times

11. Estimate the annual exceedance probability from the total of 400 000 (100 000 for each season) samples using plotting positions.

The study area and data requirements for the proposed method are described in section 2.1, while section 2.2 describes the method for determining the critical duration, and section 2.3 and 2.4 describe the generation of antecedent conditions and meteorological data series. The hydrological model is presented in section 2.5 and the method for constructing the flood frequency curve is outlined in section 2.6 and the sensitivity analysis is presented in section 2.7.

2.1 Study Area and Data requirements

2.1.1 Catchment selection and available streamflow data

The study area consists of a set in Norway, consisting of a dataset of 20 catchments located throughout Norway (fig 2) the whole country, is shown in Fig 2. All catchments have at least 10 years of hourly discharge data, and in all cases the length of the daily flow record is considerably longer than 10 years. All selected catchments are members of the Norwegian Bench Mark dataset (Fleig, 2013), which ensures that the data series are unaffected by significant streamflow regulation and have discharge data of sufficiently high quality suitable for the analyses of flood statistics. The catchment size was restricted to small and medium-sized catchments (maximum area is 854 km²), as the structure of the 3-parameter PQRUT model does not
take into account all of the storage processes within the catchment which possibly the longer-term storage processes which can contribute to delaying runoff the runoff response during storm events. Previous applications of PQRUT in Norway indicate that this shortcoming is most problematic for larger catchments. Discharge datasets with both daily and hourly time steps were obtained from the national archive of streamflow data held by NVE (https://www.nve.no/). The catchments were delineated and their geomorphological properties were extracted using the NEVINA tool: http://nevina.nve.no, except for Q, which was calculated using the available gridded data (further details are given in 2.1.2 below). In order to illustrate the application of the method, we have selected three catchments which can be considered representative for different flood regimes in Norway: Krinsvatn in western Norway, Øvrevatn in northern Norway and Hørte in southern Norway (Fig. 2).

Table 1 summarises the climatological and geomorphological properties of these three catchments, including: area \( A \) (in \( km^2 \)), mean annual runoff \( Q \) (in mm \( \text{year}^{-1} \)), mean annual precipitation \( P \) (in mm \( \text{year}^{-1} \)), mean elevation \( Hm50 \), percent of percentage forest-covered area \( (For) \), percent of percentage marsh-covered area \( (M) \), percent of percentage area with sparse vegetation above three tree line \( (B) \), ‘effective’ lake percent \( (Lk) \) and catchment steepness \( (Hl) \) and the mean annual temperature \( \text{Temp} \). The effective lake percent \( Lk \) is used to describe the ability of water bodies to attenuate peak flows such that lake areas which are closer to the catchment outlet have a higher weight than those near the catchment divide. It is calculated as \( \sum \frac{A_i \times Lk_i}{A} \times 100 \), where \( A_i \) is the area of lake \( i \), \( A^{-1} \) is the catchment area upstream of lake \( i \) and \( A \) is the total catchment area. The dominant land cover for Krinsvatn and Øvrevatn is sparse vegetation over tree line \( B \), while the land cover for Hørte is mainly forest. For Krinsvatn, the \( Lk \) is higher and the area covered by marsh, \( M \), is more than 109%. The catchment Krinsvatn steepness \( (Hl) \) (defined as \( \frac{(Hm75-Hm25)}{L} \), where \( L \) is the catchment length and \( Hm25 \) and \( Hm75 \) are the 25 and 75 quantiles of the catchment elevation) is highest for Hørte (18.7 m/km) and lowest for Krinsvatn (5.4 m/km). The catchment Krinsvatn, being located near the western coast of Norway, has a much higher mean annual precipitation \( (P) \), i.e. an average of 2354 mm \( \text{year}^{-1} \), compared to year-1, in comparison with Hørte (1261 mm/year) and Øvrevatn (1558 mm \( \text{year}^{-1} \)). The dominant flood regime for Krinsvatn is primarily rainfall-driven high flows, as the catchment is located in a coastal area and is characterised by high precipitation values and an average annual temperature of around 4°C. The highest observed floods, however, also have a contribution from snowmelt. The season of the AMAX (annual maxima flood) is the winter period, i.e. December – February, although high flows can occur throughout the year. Hørte has a mixed flood regime with most of the AMAX flood events in the period September–November, but in some years annual flood events occur in the period March–May and are associated with rainfall events during the snowmelt season. Øvrevatn has a predominantly snowmelt flood regime with most AMAX flood events occurring in the period June – August, due to the lower temperatures in the region such that precipitation falls as snow during much of the year. The catchments were delineated and their geomorphological properties were extracted using the NEVINA tool: http://nevina.nve.no, except for \( Q \), which was calculated using the available streamflow data and \( P \), which was calculated using available gridded data (further details are given in 2.1.2 below).
2.1.2 Available meteorological data

Data for temperature and precipitation with daily time resolution were obtained from seNorge.no. This dataset is derived by interpolating station data on a 1 km² grid and is corrected for wind losses and elevation (Mohr, 2008). In addition, meteorological data with a sub-daily time step is needed for calibrating the PQRUT model, as many of the catchments have fast response times. For this, precipitation and temperature data with a three-hour resolution, representing a disaggregation of the 24-hour gridded seNorge.no data using the HIRLAM hindcast series (Vormoor and Skaugen, 2013), were used. The HIRLAM atmospheric model for northern Europe has a 0.1 degree resolution (around 10 km²) and we used a temporal distribution of three hours. The HIRLAM data set was first downscaled to match the spatial resolution of the seNorge data (Vormoor and Skaugen, 2013). The precipitation of the HIRLAM data was rescaled to match the 24-hour seNorge data and (Vormoor and Skaugen, 2013). Then, these rescaled values were used to disaggregate the seNorge data to a 3-hour time resolution. The method was validated against 3-hour observations, and the correlation of the method was found to be higher than that obtained by simply dividing the seNorge data into eight equal parts hourly values (Vormoor and Skaugen, 2013). These datasets were further disaggregated to a 1-hour time step using a uniform distribution by dividing into three equal parts to match the time resolution of the discharge data, although a 3-hour time step could also be used to streamflow data.

2.1.3 Initial conditions

The stochastic PQRUT method requires time series of soil moisture deficit, SWE and initial discharge. These data series are used to generate initial conditions, which serve as input in constructing probability distribution functions for generating initial conditions for the event-based PQRUT models simulations. Sources of these data can be e.g. remotely sensed data (see, for example, the review provided in Brocca et al. (2017)) or gridded hydrological models, which can be used for modelling in ungauged basins. In this study, the DDD hydrological model was used to simulate these data series. The DDD model is a conceptual model that includes snow, soil moisture and runoff response routines and is calibrated for individual catchments using a parsimonious set of model parameters. The snowmelt routine of DDD model uses a temperature-index method and accounts for snow storage and melting for each of 10 equal area elevation zones. The soil moisture routine is based on one dynamic storage reservoir, in which we find both the saturated- and the unsaturated zone, having capacities which vary in time. The flow percolates to the saturated zone if the water content in the unsaturated zone exceeds 0.3 of its (dynamic) capacity. The response routine includes routing of the water in the saturated zone using a convolution of unit hydrographs which are based on the distribution of distances to the nearest river channel within the catchment and from the distribution of distances within the river channel.

2.2 Critical Duration

When simulating flood response with an event-based model, it is important to specify the so-called critical duration (Meynink and Cordery, 1976) to ensure that the complete flood hydrograph is modelled, flood peak is correctly modelled. The critical duration is an important quantity which effectively links the duration and the intensity of precipitation events of a given probability. In order to
determine the length of the precipitation input series producing the most extreme flows, a critical duration for storm events, defined as the duration that results in the highest observed peak value, had to be determined for each catchment. To determine the critical duration, flood events over a certain from the daily time series over a quantile threshold (in this case, 0.9) were extracted. The POT (peak over threshold) flood events were considered to be independent if they were separated by at least seven days of lower values than values below the threshold. The day with the maximum value (peak) of the peak value of streamflow was then identified for each event. The peak values were tested for correlations with the precipitation on the day of the peak flow and on days -1, -2 and -3 before the peak. The critical duration was determined as the number of days in which the correlation between the precipitation and the streamflow was higher than 0.25. This threshold value was selected because it gave realistic durations for the catchments in the study area. At first, as an alternative approach, the critical duration was also set to equal the number of days for which the correlation was significant at p=0.01. This method resulted, however, which however resulted in very long durations in some cases. A possible reason for this is that if there are only a few observations, even relatively low Pearson correlation coefficients can produce statistically significant p-values. In some catchments (mostly those having a snowmelt flood regime), no significant correlation was found between discharge and precipitation, and in this case the critical duration was fixed to 24 hours determined by considering only flood events in the September-November (SON) season in which most events are caused by rainfall (in this case, during the autumn). If the critical duration was more than one day, the precipitation was aggregated to the critical duration by applying a moving window to the data series. For Hørte and Øvrevatn, the critical duration was found to be 24 hours and for Krinsvatn it was found to be 48 hours (Fig. 3).

### 2.3 Precipitation and temperature sequence generation

In addition to the critical duration of the event, the sequence of the input data must be prescribed generated for the stochastic simulation. Snowmelt can be important in the catchments considered in this study, so both the sequence of precipitation and temperature must be considered. In order to account for seasonality, the meteorological data series were first split into standard seasons: DJF, MAM, JJA and SON. In this way, we ensure that more homogeneous samples are used to fit the statistical distributions. Although the season at risk could have been defined for each catchment individually (e.g. Paquet et al., 2013), the standard season definition was used for all catchments. Precipitation events over a threshold (POT events) were identified in the 24 h precipitation data precipitation datasets and a Generalized Pareto distribution was fitted to the series of selected events. In order to select a threshold value for event selection, two criteria were used: 1) the threshold must be higher than the 0.93 quantile, and 2) the number of selected events must be between two and three per season. Although other methods for threshold selection exist, such as the use of mean life residual plots, the described method gives adequate is much simpler to apply and gives acceptable results (e.g. Coles, 2001). The selected threshold varied between the 0.93 to 0.99 quantiles, depending on the season and catchment. In addition, storm hyetographs and the exponential distribution is often fitted to POT events, as it can give more robust results than the GP distribution. Figure 4 shows the return levels calculated from the GP and Exponential distributions and the empirical return levels and demonstrates that it is appropriate to use these models. In this case, we have preferred to use the GP due to the inclusion of the shape parameter for describing the behaviour of the highest quantiles. A exponential distribution, however, could also be used, as could a compound weather pattern-based distribution
such as the MEWP distribution (e.g. Garavaglia et al., 2010; Blanchet et al., 2015). In addition, a temperature sequence with a 1-hour time resolution were was identified from the disaggregated seNorge data, introduced in section 2.2.1.2, and extracted for each POT events. Using the fitted Generalized Pareto (GP) distribution, event The precipitation depths were simulated and the storm generated (for 100,000 events) from the fitted GP distribution for each season. Storm hyetographs were used to disaggregate the precipitation values as follows: a storm hyetograph was first sampled (fig. 9) from the extracted hyetographs for the selected POT precipitation events (by matching the dates of the selected POT precipitation events to the disaggregated seNorge dataseriesthe), taking into account seasonality, and the ratios between the 1-hour and the total precipitation for the event were calculated according to:

\[ P_{hsim} = \frac{P_i}{\text{sum}(P_{Pi})} P_{asim} \]  

where \( P_{hsim} \) is the simulated 1-hour precipitation intensity and \( P_{asim} \) is the daily intensity simulated daily intensity and \( P_i \) is the 1-hour disaggregated SeNorge intensity. The calculated ratios were then used to rescale the simulated values(fig. 9).

2.4 Antecedent snow water equivalent, streamflow and soil moisture deficit conditions

In order to determine the underlying distribution for various antecedent conditions, the relevant quantities were extracted from simulations using based on the DDD hydrological model of Skaugen and Onof (2014). The model was calibrated for the selected catchments at a daily timestep using a MCMC routine (Petzoldt, 2010). Output from DDD model runs were used to extract values for the initial streamflow, snow water equivalent (SWE) and soil moisture deficit, prior to the seasonal at the onset of the previously selected seasonal flood POT events. It is important to note that simulated values for the soil moisture deficit are used. However as described in Skaugen and Onof (2014), the model provides realistic values in comparison with measured groundwater levels. The POT event series used for this is the same as that used for identifying the critical duration (described in section 2.2).

After extracting the initial conditions, the correlation between the variables was tested for each season for each catchment. As the correlation between the variables is in most cases significant, the variables were jointly simulated using a truncated multivariate normal distribution. In order to achieve a normality for the marginal distributions, the SWE and the discharge were log-transformed. In the spring and summer, the SWE is often very low or 0 in some catchments. If the proportion of non-zero values, \( p_p \), was greater than 0.3 (around 15 observations), the values were simulated using a mixed distribution as:

\[ F(x) = pG_1(x) + (1 - p)G_2(x) \]  

where \( G_1 \) and \( G_2 \) represent the multivariate normal distribution with discharge, soil moisture deficit and SWE as variables and (denoted as \( x \)) and \( G_2 \) the bivariate normal distribution for the discharge and soil moisture, respectively. The probability \( p \) for switching between the trivariate and bivariate distributions is based on the historical data for SWE higher than 0. In addition, because the initial conditions are not expected to include extreme values, the values of the initial conditions were truncated to be between the minimum and maximum of the observed ranges. The correlation between the observed and
simulated variables is shown in Figure 5 for the Krinsvatn catchment, and although the distribution of simulated values exhibits a very good resemblance to that of the observed, there is not a perfect correspondence between the two. A reason for this may be that the variables (even after log transformation) do not exactly follow a normal distribution. We considered using copulas for the correlation structure of the initial conditions \cite{Hao2016} however \cite{Hao2016}. However, as the data were limited \cite{Filipova2016} (around 50 observations per season), these were much more difficult to fit. Similarly, nonparametric methods such as kernel density estimation were not deemed to be feasible due to the limited number of observations. Therefore, the multivariate normal distribution was chosen as the best alternative for modelling the joint dependency between the variables comprising the initial conditions for the stochastic modelling.

### 2.5 PQRUT model

The PQRUT \cite{PQRUT2007} model was used to simulate the streamflow for the selected storm events. The PQRUT model is a simple, event-based, 3-parameter model \cite{Filipova2016} which is used, amongst other things, for estimating design floods and safety check floods for dams in Norway \cite{Wilson2011}. In practical applications, a hypothetical precipitation design sequence of a given return period is routed through the PQRUT model, usually under the assumption of full catchment saturation. For this reason, only the hydrograph response is simulated, and there is no simulation of subsurface and other storage components, such as are found in more complex conceptual hydrological models. Of the three model parameters, $K_1$ corresponds to the fast hydrograph response of the catchment, and the parameter $K_2$ is the slower or ‘delayed’ hydrograph response. The parameter $T_{rt}$ is the threshold above which $K_1$ becomes active.

The PQRUT model was calibrated for flood events for each catchment the 45 highest flood events by using the DDS (Dynamically Dimensioned Search) optimization \cite{Tolson2007} routine \cite{Tolson2007} and the Kling Gupta efficiency (KGE) criterion \cite{Gupta2009} was used as the objective function. The general procedures used for the PQRUT calibration are described in Filipova et al. \cite{Filipova2016}. As described in that work, an additional parameter, the soil deficit, $L_p$, was introduced to account for initial losses to the soil zone and is necessary if one is to achieve calibration of the model to actual events (rather than hypothetical events in which the catchment is fully saturated). The reason for this is that, even though fully saturated conditions are assumed when the model is used to estimate PMF or other extreme floods with low probabilities, the model needs to account for initial losses when actual (more frequent) events are simulated. This procedure is described in more detail in Filipova et al. \cite{Filipova2016}. In addition, regional values can be used in ungauged or poorly catchments \cite{Andersen1983, Filipova2016}.

For the work presented here, the value of this parameter $L_p$ was set to the initial soil moisture deficit, estimated using DDD. This parameter variable functions as an initial loss to the system, such that the input precipitation to the reservoir model is 0 until the value of $L_p$ is exceeded by the cumulative input rainfall. In order to model flood events involving snowmelt, a simple temperature index snow melting rate was used:

$$S = C_s(T - T_L)$$

\cite{Wilson2011}. 

10

15

20

25

30
where $S$ is the snow melting rate in mm/hour, $Cs$ is a coefficient accounting for the relation between temperature and snowmelt properties and $T_L$ is the temperature threshold for snowmelt (here fixed at 0°C). The model used regional [Regional] values for the Cs parameters related to the as a function of catchment properties, based on the ranges given in Midtønnmne and Pettersson (2011) were applied. In addition, the temperature threshold between rain and snow was set to $T_X = 0.5^0C$, which is typically used in Norway (Skaugen, 1998).

2.6 Flood frequency curves

Seasonal and annual flood frequency curves were constructed by extracting the peak discharge for each event and estimating the plotting positions of the points using the Gringorten plotting position formula:

$$P_e = \frac{(m - 0.44)}{(N + 0.12)k}$$

where $P_e$ is the exceedance probability of the peak, $m$ is the rank (sorted in decreasing order) of the peak value, $N$ is the number of years, $k$ is the number of events per year. The number of events per year, $k$, was set to be equal to the average number of extracted POT storm (precipitation) events per year. These simulated events were compared with the POT flood events extracted from the observations (Fig. 7). After calculating the probability of the simulated events using Eq. 4, the initial conditions and seasonality for a return period of interest can be extracted. For example, the events with return period between 90 and 110 years were extracted (representing around 80 events), and the hydrological conditions for those events were identified (Table 2). The results show that there is a large variation in precipitation values, total precipitation depths and initial conditions that can produce flood events of a given magnitude and this is the reason why it is difficult to assign initial conditions in event-based models. However, it is still useful to extract the distribution of these values in order to ensure that the ranges are reasonable and the catchment processes are properly simulated. For example, the average snowmelt is negative (i.e., there is snow accumulation) for Krinsvatn, which means that in most cases snowmelt does not contribute to the extreme floods. This is reasonable as the catchment is located in western Norway, where the climate is warmer (the mean temperature is around 4°C) and the mean elevation is low. The average snowmelt contribution for Øvrevatn is much higher as this catchment has a predominantly snowmelt flood regime. The soil moisture deficit for the three catchments is larger than 0, even though fully-saturated conditions are used in the event-based PQRUT model, floods with relatively long return periods (i.e., between 90 and 100 years) are being sampled here. The seasonality of the simulated values is consistent with the seasonality of the observed annual maxima (Table 1).

2.7 Sensitivity analysis

A sensitivity analysis was performed for the three test catchments, Hørte, Øvrevatn and Krinsvatn, in order to determine the relative importance of the initial conditions, precipitation and the parameters of PQRUT, the effect of the random seed and length of simulation on the flood frequency curve. To test the sensitivity of the model, we have used several different model runs and calculated the percentage difference of each of these model runs relative to the standard model setup, as shown in
More detailed information on the set up is given in Table 3. As these catchments are located in different regions and exhibit different climatic and geomorphic characteristics, we hypothesize that the flood frequency curve will be sensitive to different parameters, and hydrological states, precipitation, snow and catchment characteristics, as well as local climate and catchment characteristics. The results are presented in figure 8 and summarised in Table 4.

Considering the effect of the initial conditions, using fully saturated conditions results in the slight overestimation for all catchments of flood values, as expected, and the impact is higher at lower return periods. In addition, Øvrevatn shows higher sensitivity (around 30% for Q1000) to the initial soil moisture conditions than the other two catchments. A possible explanation is that the baseflow index (BFI) is higher (BFI=0.6) for Øvrevatn, than the BFI for Krinsvatn (BFI=0.4) and Hørte (BFI=0.5). This indicates that the runoff at Øvrevatn is less responsive to the rainfall input. Similarly, a sensitivity analysis by Svensson et al. (2013) shows that high sensitivity of floods to soil moisture deficit is more present for permeable catchments, where the BFI is also high. The initial discharge value does not seem to have a large impact for any of the catchments. This means that in ungauged catchments the median value can be used. If no snow component (no snowmelt and no snow accumulation) is used, there is not much difference in the results for Øvrevatn and Hørte, but the seasonality of the flood events is changed. For example, the season when the Q1000 is simulated for Øvrevatn is SON instead of JJA when most of the AMAX values are observed. Due to the change of seasonality, the precipitation values that produce Q1000 are higher (the median is around 30% higher). The soil moisture deficit, as expected, is also somewhat higher and shows much more spread with values up to 60 mm. In addition, Krinsvatn shows high sensitivity to snowmelt (29% higher) and also a step change in the frequency curve, even though the soil moisture deficit is higher. This can also be explained by the fact that the snowmelt contribution is negative, as can also be seen in Table 2.

The results for the sensitivity to the rainfall model are presented in Fig 8a. The results show that the temporal patterns of the rainfall input have high a large impact (up to 50%) on the flood frequency curve for Hørte and Krinsvatn, as these catchments have a predominantly rainfall-dominated flood regime, but the impact is very little for Øvrevatn. High A high sensitivity to the shape of the hyetograph was also found in Alfieri et al. (2008). They found that using rectangular hyetograph results in a significant underestimation of the flood peak while the Chicago hyetograph (e.g. Chow et al., 1988) resulted in overestimation. In addition, Øvrevatn and Hørte showed sensitivity (28-9 around 20%) to the choice of the statistical distribution for modelling precipitation. This means that the uncertainty in fitting the rainfall model can propagate to the final results of the stochastic PQRUT, and therefore, it is important to ensure that the choice of distribution and parameters should be carefully considered. A high sensitivity to the parameters of the rainfall model was also described by Svensson et al. (2013), who suggest that this is a main source of uncertainty. Both Hørte and Krinsvatn showed relatively lower sensitivity to the threshold value for the GP distribution, compared to Øvrevatn. A reason for the high sensitivity to the threshold value for Øvrevatn is that using a higher quantile for the threshold leads to selecting fewer events, which are less representative and the fit of the GP becomes more uncertain leading to a higher degree of uncertainty in the GP fit.

In general addition, all catchments are very sensitive to the parameters of the PQRUT model (Fig 8b) and there is a high large uncertainty in these values. Because of the higher sensitivity to the calibration of the rainfall-runoff model, a conclusion can be made that, in practice, if streamflow data is available it is important that this is used for calibrating the PQRUT model.
Considering the effect of the initial conditions (Fig 8c), using fully saturated conditions results in the slight overestimation for all catchments of flood values, as expected, and the impact is higher at lower return periods. In addition, Øvrevatn shows a higher sensitivity (around 26% for Q1000) to the initial soil moisture conditions than the other two catchments. A reason for this is that for Øvrevatn, higher soil moisture conditions are associated with higher rainfall quantiles. For example, for Øvrevatn, precipitation depths with a 1000-year return period are associated with median soil moisture conditions of 37 mm, while for Krinsvatn, it is 30.8 mm and for Hørte, it is 16.7 mm. The initial discharge value does not seem to have a large impact for any of the catchments. If no snow component (no snowmelt and no snow accumulation) is used, there is not much difference in the results for Øvrevatn and Hørte, but the seasonality of the flood events is changed. For example, the season when the Q1000 is simulated for Øvrevatn is SON instead of JJA when most of the AMAX values are observed. Due to this change of seasonality, the precipitation values that produce Q1000 are accordingly higher (the median is around 15% higher). The soil moisture deficit, as expected, is also somewhat higher and shows much more spread, with values up to 45 mm. In addition, Krinsvatn shows a high sensitivity to the snowmelt component (21% higher) and also a step change in the frequency curve, even though the soil moisture deficit is higher. This can also be explained by the fact that the snowmelt contribution is negative (there is snow accumulation), as can also be seen in Table 2. Other studies have also shown that the soil saturation level is not as important in comparison with the parameters of the hydrological model. For example, Brigode et al. (2014) tested the sensitivity of the SCHADDEX model using a set of block bootstrap method. In each of these experiments, different sub-periods selected from the observation record were used in turn to calibrate the rainfall model, the hydrological model and to determine the sensitivity to the soil saturation level. The results showed that for extreme floods (1000–year return period), the model is sensitive to the calibration of the rainfall and the hydrological models, but not so much to the initial conditions.

The stochastic PQRUT model shows some sensitivity to the random seeds (Fig 8d), especially for higher return periods. This is expected as the higher quantiles are calculated using a smaller sample of simulated events. Similarly, the effect of the simulation length has a larger impact on the higher quantiles (e.g., Q1000). However, the length of the simulation will depend on the required return level, as shorter simulation length can be acceptable for lower return periods, e.g., Q100.

3 Comparison with standard methods

3.1 Implementation of the methods and results

The results of the stochastic PQRUT method for the 100- and 1000-year return level were compared with the results for statistical flood frequency analysis and with the standard implementation of the event-based PQRUT method (in which full saturation and snow melting rates are assumed a priori) for the twenty test catchments described in section 2.1. For the statistical flood frequency analysis, the annual maximum series were extracted from the observed daily mean streamflow series. The GEV distribution was fitted to the extracted values using the L-moments method and the return levels were estimated. In order to obtain instantaneous peak values, the return values were multiplied by empirical ratios, obtained from regression equations, as given in (Midtømme and Pettersson, 2011). The ratios can vary substantially from catchment to catchment, and in this study, the values are from 1.02 to 1.82, depending on the area and the flood generation.
process *(snowmelt or precipitation)*. Although much more sophisticated methods could be used to obtain statistically-based return levels, the procedure used here is equivalent to that currently used in standard practice in design flood analysis in Norway. In addition, a study by Kobierska et al. (2017) showed that the GEV along with the G (Generalised logistic) distribution gives the most reliable results based on a sample of 280 catchments in Norway—the length of the daily streamflow series justifies the use of at-site flood frequency analysis (Kobierska et al., 2017); the minimum length is 31 years, while the median is 65 years of data. However, it is expected that the uncertainty will be high when the fitted GEV distribution is extrapolated to a 1000-year return period. The 1000-year return period is used here, however, as it is required for dam safety analyses in Norway (e.g. Midttømme, et al., 2011; Table 1). More robust, but potentially less reliable, estimates could be obtained using a 2-parameter Gumbel, rather than a 3-parameter GEV distribution (Kobierska et al., 2017). The standard implementation of PQRUT involves using a precipitation sequence that combines different intensities, obtained from growth curves based on the 5-year return period value *fitted using a Gumbel distribution while the ratios between the different durations are derived from empirical distribution* (Førland, 1992). The precipitation intensities were combined to form a single symmetrical storm profile with the highest intensity in the middle of the storm event. Here, whereas the storm profile is randomly sampled for the stochastic PQRUT model (Fig 9). In the application reported here, the duration of the storm event was assumed to be the same as that used for the stochastic PQRUT model. The initial discharge values were similarly fixed to the seasonal mean values, *as is common in standard practice*. The snowmelt contribution for the 1000-year return period was, in this case, assumed to be 30 mm/day for all catchments, which corresponds to 70% of the maximum snowmelt, estimated as 45 mm/day by using temperature-index factor of 4.5 mm/°C day and 10° C. The snowmelt contribution for the 100-year return period was assumed to be 21 mm/day for all catchments. In addition, fully saturated conditions were assumed for both the estimation of the 100- and 1000-year return periods. A similar implementation of PQRUT for the purpose of comparing different methods has also been described in (Lawrence et al., 2014). The parameters of the PQRUT were estimated by using the regional equations derived in Andersen et al. (1983), as these are still used in standard practice.

The performance of the three models was validated by using two different tests. Test 1 assessed whether the estimated values for the flood frequency curve 100-year return period are within the confidence intervals of a GP distribution fitted to the streamflow data with a 1-hour time step. The stochastic PQRUT shows good agreement with the observations (Fig. 10), and for 45 of 20 catchments, all the points of the derived flood frequency curve were inside the confidence intervals. As expected, for most of the catchments (44 out of 20) the return levels calculated using statistical flood frequency analysis based on the GEV distributions using daily values were within the confidence intervals. As discussed, due to the difficulty in assigning initial conditions for the event-based For the standard PQRUT model, it was only possible to estimate values for the 100- and 1000-year return periods. For this model, the values of the 100-year return level were within the confidence interval for only six of the catchments when the regional equations for the PQRUT model were used and for only eight of the catchments when calibrated parameters are used. In addition, test 2, the results of the flood frequency analysis and the Stochastic PQRUT methods were compared, based on a quantile score (QS) suggested by E. Paquet (personal communication), this is
given in Eq 5:

\[
Q_{\text{score}} = 1 - \sum (\text{abs}(Q_{\text{mod}_i} - Q_{\text{obs}_i})(Q_{\text{obs}_i} - Q_{\text{obs}_{i-1}}))
\]  

(5)

In Eq. 5–5 the observed probabilities (Qobs_i) are calculated using Gringorten positions for the POT series—peak AMAX series that were derived from the daily values. The modelled probabilities that correspond to the observed events are calculated by using the statistical flood frequency analysis and the Stochastic PQRUT model, as described previously. The standard implementation of the event-based PQRUT model was not evaluated based on QS as initial conditions could not be assigned for low return periods. As this model is usually used to calculate high quantiles (Q100 or higher), fully saturated conditions are assumed for its implementation. The results for the quantile score show similar performance, the median is around 0.83–0.65 for both methods. However, the results vary between catchments as shown in Fig. 11. Although it is difficult to evaluate the performance of the models when the dataseries are relatively short, based on the results of test 1, we can conclude that the performance of the standard PQRUT model is poorer than the performance of the statistical flood frequency analysis and the stochastic PQRUT model for the selected catchments, while the results of test 2 indicate that both the GEV distribution and the stochastic PQRUT provide similar fits to observed quantiles.

### 3.2 Discussion

A comparison of the three methods—stochastic PQRUT with the standard methods for flood estimation shows that there is a large difference between the results of the different methods (fig-three methods for both Q100 and Q1000 (Fig. 12 and 13). The violin plots (fig-boxplots (Fig. 12) show that the stochastic PQRUT method gives slightly lower results on average than the standard PQRUT model for Q100 and Q1000. This is probably due to assuming fully saturated conditions when applying the standard PQRUT for Q100, which might not be realistic for some catchments. For example, the results for the initial conditions for the three catchments, presented in section 2.6, show that the soil moisture deficit is larger than 0. However, when the results are compared for 0 for Q1000, the stochastic PQRUT gives slightly higher results. Reasons for this may be that higher precipitation intensity or snowmelt is used. Q100. Furthermore, the absolute differences between the two methods are larger in catchments with lower temperature (fig-Fig. 12). This indicates that the performance of the standard PQRUT model is worse in catchments with a snowmelt flood regime, which might be due to the difficulty in determining snowmelt contribution or to the poorer performance of the regional parameters in catchments with a snowmelt flow regime. Although showing the same—it shows a similar pattern, the standard PQRUT model, implemented using calibrated parameters results in much less spread than the implementation using the regionalised parameters, when compared to both the GEV distribution and the stochastic PQRUT model. This means that the hydrological model can introduce a large amount of uncertainty, as also indicated by the sensitivity analysis described in section 2.7 and previous results presented by Brigo de et al. (2014).

The differences between the stochastic PQRUT model and GEV is the GEV fits are much smaller than the differences between the standard PQRUT and GEV model and the GEV fits, even when calibrated parameters are used. In general, the stochastic PQRUT model gives higher values than the GEV distribution, which might be due to the uncertainty in estimating the parameters for the GEV distribution. For example, the study by Rogger et al. (2012) shows that the
flood frequency analysis based on fitting a Gumbel distribution to AMAX series underestimates high flows in catchments with a high storage capacity, where a step change in the flood frequency curve occurs. The results of the study by Rogger et al. (2012) can be explained by the fact that the Gumbel distribution, which has a shape parameter of 0 and so is not as flexible as the GEV distribution. In this study we find very low correlation between the difference between the stochastic PQRST and the GEV distribution and catchments with high values of the effective lake index or the percentage of the catchment covered by marsh (fig 13), where we would expect that the storage capacity is higher. In addition, the for the PQRST modelling. The differences are larger (i.e. the stochastic PQRST results are lower, as shown in fig 13) in western Norway where P and Q are higher and for catchments with higher catchment steepness HI (defined as (Hm75-Hm25)/L, where L is the catchment length and Hm25 and Hm75 are the 25 and 75 quantiles of the catchment elevation) steeper catchments, i.e. with a higher value of HI. A reason for this might be that the empirical ratios that are used to convert daily to peak flows in these catchments are inaccurate and possibly too high.

Similarly to the violin-plots, fig 13 boxplots, Fig. 13 also shows that the results of the stochastic PQRST closely match the GEV distribution fits with differences within 50% for most locations. There is no clear spatial pattern in the differences between estimates based on the GEV distribution and on the standard PQRST model, except for the catchments in mid-Norway, i.e. Trøndelag (including catchment Krinsvatn), where the GEV distribution produces higher results. However, a much larger sample of catchments is needed to assess whether there is a spatial pattern in the performance of the methods.

4 Conclusions

In this article, we have presented a stochastic method for flood frequency analysis based on a Monte Carlo simulation to generate rainfall hyetographs and temperature series to drive a snowmelt estimation, along with the corresponding initial conditions. A simple rainfall-runoff model is used to simulate discharge, and plotting positions are used to calculate the final probabilities.

In this way, we can generate thousands of flood events and use the empirical distribution instead of extrapolating a statistical distribution fitted to the observed events. The approach thereby gives significant insights into the various combinations of factors that can produce floods with long return periods in a given catchment, including combinations of factors that are not necessarily well represented in observed flow series. It is thus a very useful complement to statistical flood frequency analysis and can be particularly beneficial in catchments with shorter streamflow series compared to the precipitation record as well as in ungauged catchments.

In order to apply the method, we assume that the precipitation and temperature series are not significantly correlated with the initial conditions, which allows us to simulate them as independent variables. Although we have not performed a statistical analysis, the independence between the flood precipitation events and the initial conditions has been verified by e.g. Paquet et al. (2013)(Paquet et al., 2013). Due to the considerable seasonal variation in the initial conditions, seasonal distributions were used. In addition to obtaining more homogeneous samples, this allows one to check for a check of the seasonality of the flood events, which can be of interest in catchments with a mixed flood regime. In this study, we have used a GP distribution to model the extreme precipitation. However, if only shorter precipitation datasets are available, the exponential distribution
or even regional frequency analysis methods may provide more robust results. A limitation of the method is that PQRUT can only be used for small and medium-sized catchments, since its three parameters cannot take into account the spatial variation in the snowmelt and soil saturation conditions within the catchment. However, for the catchments presented in this study (all with a catchment area under 850 km²), the model produces relatively good fits to the observed peaks, even though it uses a very limited number of parameters. For example, a semi-distributed temperature-index snowmelt model, such as the one used in HBV (Sælthun, 1996), may improve the results in some catchments, though this would also increase the amount of data required. 

In this study, initial conditions based on simulations using a hydrological model (DDD) were used. However, in other applications, initial conditions may be based on remotely-sensed data, or on the output of gridded hydrological models. This is particularly important for the application of the method in ungauged basins. This requires that this model is calibrated at each catchment. Considering the results of the sensitivity analysis, the quality of the initial conditions is not as important as that of the precipitation data for the estimation of extreme floods (with return periods higher than 100 years). This means that if no other data is available, the output of gridded hydrological model could be considered as a source of this input data. Alternatively, remotely sensed data can be used for soil moisture and the snow water equivalent while regional values for the initial discharge can be derived. This for example, can be an option in ungauged basins.

The stochastic PQRUT model was applied to 20 catchments, located in different regions of Norway and was compared with the results of the statistical flood frequency analysis and the event-based PQRUT model. This comparison shows that there are large differences between the methods. Major sources of uncertainty for the flood frequency analysis are the use of short data series and the empirical peak to volume ratios that were used to calculate the instantaneous flow. There is also uncertainty in the event-based rainfall-runoff simulation method because of difficulties in assigning the initial conditions and in calibrating the rainfall-runoff model. This first of these is a reason why the use of a stochastic model is important, as it can simulate multiple initial conditions and easily incorporate the uncertainty associated with this choice. Event-based PQRUT method which is today used in standard practice. Due to the high uncertainty in estimating extreme floods, the application of different methods most often the different methods produces differing results, as in often the case in practical applications. However, in this work we have shown that the stochastic PQRUT model gives estimates which generally are more similar to those obtained using a statistical flood frequency analysis based on the observed annual maximum series than are estimates obtained using a standard implementation of PQRUT. As it is not possible to test the reliability of estimates for the 500- or 1000-year flood (due to length of the observed streamflow series relative to the return period of interest), the use of alternative methods for flood estimation, including stochastic simulations such as presented here, is an essential component of flood estimation in practice. A possible way forward is to consider estimates based on different methods by calculating a weighted average of the various estimates, in which the weighting is based on an assessment of the uncertainty characterizing the individual methods.

Code and data availability. The R package StochasticPQRUT (https://github.com/valeriyaafilipova/StochasticPQRUT) can be installed from github and contains sample data.
Competing interests. On behalf of all authors, the corresponding author states that there is no conflict of interest.

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References


**Figure 1.** Diagram of the Stochastic PQRUT model
Figure 2. Location of the selected catchments. The catchments Hørte, Øvrevatn and Krinsvatn, for which we show the method in more detail, are plotted in red.
Figure 3. Critical diagram representing the process used to establish the critical duration for Krinsvatn. The stars represent the degree of significant correlation between $Q_{obs}$ and $P$ on the day of the peak and -1, -2 and -3 days before the peak at $p=0.01$. The critical duration is set in this case to two days because the correlation is over 0.25 between $Q_{obs}$, $Q_{obs}$ vs $P$ and $P_1$ is greater than 0.25.
Figure 4. Storm pattern scaling—Return level plots for the simulated events, sampled event is shown in dark grey fitted Generalised Pareto (GP) and exponential (EXP) distributions for peak over threshold precipitation events (GP – red; EXP – green) for the simulated storm event in light-grey three selected catchments.
Figure 5. Correlation scatterplots for initial condition conditions (Snow water equivalent (SWE), soil moisture deficit (lp), initial discharge (Qobs)) for the Krinsvatn catchment for POT events (flood events over 0.9 quantile, 57 observations). Scatterplots for observed quantities are shown on the left and simulated on the right (based on 100 000 simulations). The stars represent the degree of significant correlation.
Figure 6. Structure of the PQRUT model
Figure 7. Comparison between the observed and simulated flood frequency curves for Horte, Krinsvatn and Øvrevatn.
Figure 8. Sensitivity analysis of the Stochastic PQRUT for the following variables: a) the rainfall model, b) the effect of the simulation length and the parameters of the hydrological model, c) the initial discharge values, snowmelt conditions and catchment saturation, and the d) the random seed used for the simulations. Detailed information on the set up is given in table 3.
Figure 9. Storm patterns used for the simulated events in the stochastic PQRUT model (left) where violin plots are used to show the density of the intensities and the storm pattern typically used with the standard PQRUT model (right). P represents the ratio of the hourly precipitation to the total precipitation depth for the event.
Figure 10. Number of models giving estimates for Q100 within the confidence intervals of the GP distribution fitted to 1-hour streamflow data.
Figure 11. Boxplots of Quantile score, calculated (Eqn. 5) for the Flood estimates based on statistical flood frequency analysis using a GEV distribution (red) and Stochastic on the stochastic PQRUT model. Catchments are represented by different colours (green).
Figure 12. Violin and boxplots showing the distribution of the differences (calculated by dividing the estimates by the average of all of the models) between the Stochastic PQRUT, PQRUT and GEV for the 100- and 1000–year return level.
Figure 13. Results of the comparison between stochastic PQRUT, PQRUT and GEV for the values of the 1000-year return level. The absolute differences (calculated by dividing the estimate by the average of all models) are correlated with catchment properties. Positive correlations are given by red colour represents positive correlation and negative by blue colour negative correlation.
Table 1. Properties for the selected catchments, including Area-catchment area, $Q$-mean annual streamflow, $P$-mean annual precipitation, $M$-percent marsh, $B$-percent sparse vegetation over treeline, $Hl$-catchment steepness, $Lk$-effective lake percent and $Temp$-mean annual temperature.

<table>
<thead>
<tr>
<th>Station</th>
<th>Area, km²</th>
<th>$Q$, mm/year</th>
<th>$P$, mm/year</th>
<th>Hm50, m</th>
<th>For, %</th>
<th>M, %</th>
<th>B, %</th>
<th>Lk, %</th>
<th>Temp, Hl/m/km</th>
<th>Temp, °C</th>
<th>Season of AMAX</th>
</tr>
</thead>
<tbody>
<tr>
<td>Hørte</td>
<td>157</td>
<td>961</td>
<td>1261</td>
<td>501</td>
<td>73</td>
<td>3</td>
<td>18</td>
<td>0.3</td>
<td>18.7</td>
<td>2.89</td>
<td>SON</td>
</tr>
<tr>
<td>Krinsvatn</td>
<td>207</td>
<td>1890</td>
<td>2354</td>
<td>348</td>
<td>20</td>
<td>9</td>
<td>57</td>
<td>1.1</td>
<td>5.4</td>
<td>4</td>
<td>DJF</td>
</tr>
<tr>
<td>Øvrevatn</td>
<td>526</td>
<td>1448</td>
<td>1558</td>
<td>564</td>
<td>35.2</td>
<td>2.5</td>
<td>52</td>
<td>0.6</td>
<td>14.8</td>
<td>-0.14</td>
<td>JJA</td>
</tr>
</tbody>
</table>
Table 2. Precipitation and initial conditions for \(Q_{100}\), range corresponds to 5 and 95 percentile. For the seasonality, the actual fraction of the simulated events is given in brackets.

<table>
<thead>
<tr>
<th>Station name</th>
<th>P mean, mm</th>
<th>P median, mm</th>
<th>P range, mm</th>
<th>Q mean, m(^3)/s</th>
<th>Q median, m(^3)/s</th>
<th>Q range, m(^3)/s</th>
<th>Snowmelt mean, mm</th>
<th>Snowmelt median, mm</th>
<th>Snowmelt range, mm</th>
<th>Soil Moisture mean, mm</th>
<th>Soil Moisture median, mm</th>
<th>Soil Moisture range, mm</th>
<th>Season</th>
</tr>
</thead>
<tbody>
<tr>
<td>Krinsvatn</td>
<td>157.4</td>
<td>143.2</td>
<td>140.4</td>
<td>140.1-248.9</td>
<td>19.5</td>
<td>13.6</td>
<td>4 - 42.8</td>
<td>-5.3-1.7</td>
<td>0</td>
<td>-29.3-9.6</td>
<td>110.1-248.9</td>
<td>107-217</td>
<td>SON (0.5)</td>
</tr>
<tr>
<td>Hørte</td>
<td>70</td>
<td>66.1</td>
<td>67</td>
<td>51.94-100.5</td>
<td>11.3</td>
<td>9.4-10</td>
<td>192.6-27.1</td>
<td>6.8-4.9</td>
<td>0</td>
<td>0-37 - 22.7</td>
<td>1.9-21.6</td>
<td>0-23.2</td>
<td>DJF (0.4)</td>
</tr>
<tr>
<td>Ørevatn</td>
<td>55.73</td>
<td>52.5</td>
<td>55.9</td>
<td>35-95-38-127</td>
<td>66.5</td>
<td>62.6</td>
<td>23-106.2</td>
<td>29-25.1</td>
<td>35.8-30</td>
<td>-6.11-42.9</td>
<td>2.6-44.2</td>
<td>0-49.8</td>
<td>JJA (0.76)</td>
</tr>
</tbody>
</table>

For the seasonality, the actual fraction of the simulated events is given in brackets.
<table>
<thead>
<tr>
<th>Variable</th>
<th>Standard model setup</th>
<th>Model set up for sensitivity analysis</th>
</tr>
</thead>
<tbody>
<tr>
<td>Soil moisture deficit</td>
<td>generated from multivariate normal distribution</td>
<td>Fully saturated conditions ( lp=0 ) (no initial loss)</td>
</tr>
<tr>
<td>Snowmelt</td>
<td>snowmelt component (described in section 2.5)</td>
<td>No snowmelt component</td>
</tr>
<tr>
<td>Initial discharge</td>
<td>generated from multivariate normal distribution</td>
<td>use median discharge for each season</td>
</tr>
<tr>
<td>Precipitation intensity</td>
<td>disaggregation using random historic storm events</td>
<td>divide into 24 equal parts</td>
</tr>
<tr>
<td>Precipitation depth-distribution</td>
<td>generated from GP distribution</td>
<td>Exponential distribution</td>
</tr>
<tr>
<td>Precipitation depth-threshold</td>
<td>threshold is selected between 0.93-0.99 quantile</td>
<td>GP fitted to 0.99 quantile</td>
</tr>
<tr>
<td>parameters of PQRUT</td>
<td>calibrated to selected storm events</td>
<td>use Latin Hypercube to sample 50 values within the 5% and 95% confidence intervals for the regression equations for ( K_1, K_2 ) and ( Trt ).</td>
</tr>
<tr>
<td>length of simulation</td>
<td>400,000</td>
<td>Sequence of lengths were used staring at 40,000 to 400,000 by increment of 40,000</td>
</tr>
<tr>
<td>effect of random seed</td>
<td>random seed is set at the start</td>
<td>50 different runs using different random seeds</td>
</tr>
</tbody>
</table>
Table 4. Percent difference between the model runs of the sensitivity analysis to the calibrated model

<table>
<thead>
<tr>
<th>Catchments</th>
<th>Hørte</th>
<th>Krinsvatn</th>
<th>Øvrevatn</th>
<th>Hørte</th>
<th>Krinsvatn</th>
<th>Øvrevatn</th>
<th>Hørte</th>
<th>Krinsvatn</th>
<th>Øvrevatn</th>
</tr>
</thead>
<tbody>
<tr>
<td>setup/return period</td>
<td>Q10</td>
<td>Q10</td>
<td>Q10</td>
<td>Q100</td>
<td>Q100</td>
<td>Q100</td>
<td>Q100</td>
<td>Q100</td>
<td>Q100</td>
</tr>
<tr>
<td>1. 100–50 values, sampled using latin hypercube within the 5% and 95% confidence intervals for the regression equations for K1,K2 and T max</td>
<td>1.2%–36.7%</td>
<td>1.2%–35.3%</td>
<td>1.6%–36.7%</td>
<td>1.2%–35.3%</td>
<td>1.6%–36.7%</td>
<td>1.2%–35.3%</td>
<td>1.6%–36.7%</td>
<td>1.2%–35.3%</td>
<td>1.6%–36.7%</td>
</tr>
<tr>
<td>2. 100–50 values, sampled using latin hypercube within the 5% and 95% confidence intervals for the regression equations for K1,K2 and T min</td>
<td>1.2%–36.7%</td>
<td>1.2%–35.3%</td>
<td>1.6%–36.7%</td>
<td>1.2%–35.3%</td>
<td>1.6%–36.7%</td>
<td>1.2%–35.3%</td>
<td>1.6%–36.7%</td>
<td>1.2%–35.3%</td>
<td>1.6%–36.7%</td>
</tr>
<tr>
<td>3. GPD was fitted to 0.99 quantile</td>
<td>0.1%–10.0%</td>
<td>0.1%–13.5%</td>
<td>0.1%–16.5%</td>
<td>0.1%–14.4%</td>
<td>0.1%–16.5%</td>
<td>0.1%–14.4%</td>
<td>0.1%–16.5%</td>
<td>0.1%–14.4%</td>
<td>0.1%–16.5%</td>
</tr>
<tr>
<td>4. Exponential distribution instead of GP distribution</td>
<td>0.3%–3.0%</td>
<td>0.3%–3.0%</td>
<td>0.3%–3.0%</td>
<td>0.3%–3.0%</td>
<td>0.3%–3.0%</td>
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<td>0.3%–3.0%</td>
<td>0.3%–3.0%</td>
</tr>
<tr>
<td>5. Disaggregate precipitation depth using uniform distribution (constant intensity) instead of using temporal patterns</td>
<td>-2.5%–7.3%</td>
<td>-2.5%–7.3%</td>
<td>-2.5%–7.3%</td>
<td>-2.5%–7.3%</td>
<td>-2.5%–7.3%</td>
<td>-2.5%–7.3%</td>
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<td>-2.5%–7.3%</td>
<td>-2.5%–7.3%</td>
</tr>
<tr>
<td>6. Median discharge instead of randomly generated</td>
<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
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<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
</tr>
<tr>
<td>7. No snowmelt modelled</td>
<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
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<td>-2.4%–2.3%</td>
<td>-2.4%–2.3%</td>
</tr>
<tr>
<td>8. Fully saturated conditions</td>
<td>20.6%–20.9%</td>
<td>20.6%–20.9%</td>
<td>20.6%–20.9%</td>
<td>20.6%–20.9%</td>
<td>20.6%–20.9%</td>
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<td>20.6%–20.9%</td>
<td>20.6%–20.9%</td>
<td>20.6%–20.9%</td>
</tr>
<tr>
<td>Different simulation length 40 000 to 400 000 simulations by 40 000 simulations (range)</td>
<td>4.9%–7.3%</td>
<td>4.9%–7.3%</td>
<td>4.9%–7.3%</td>
<td>4.9%–7.3%</td>
<td>4.9%–7.3%</td>
<td>4.9%–7.3%</td>
<td>4.9%–7.3%</td>
<td>4.9%–7.3%</td>
<td>4.9%–7.3%</td>
</tr>
<tr>
<td>50 simulations with different random seeds (range)</td>
<td>-0.6%–0.8%</td>
<td>-0.6%–0.8%</td>
<td>-0.6%–0.8%</td>
<td>-0.6%–0.8%</td>
<td>-0.6%–0.8%</td>
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<td>-0.6%–0.8%</td>
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