Towards predictive data-driven simulations of wildfire spread – Part 2: Ensemble Kalman Filter for the state estimation of a front-tracking simulator of wildfire spread

M. C. Rochoux¹,²,³, C. Emery¹,²,⁴, S. Ricci¹,², B. Cuenot¹, and A. Trouvé⁴

¹CERFACS, 42 avenue Gaspard Coriolis, 31057 Toulouse Cedex 01, France
²SUC/CNRS-URA1875, 42 avenue Gaspard Coriolis, 31057 Toulouse Cedex 01, France
³Ecole Centrale Paris, CNRS-UPR288, Grande voie des vignes, 92295 Châtenay-Malabry, France
⁴Dept. of Fire Protection Engineering, University of Maryland, College Park, MD 20742, USA

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Correspondence to: M. C. Rochoux (melanie.rochoux@graduates.centraliens.net)

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Abstract

This paper is the second part in a series of two articles, which aims at presenting a data-driven modeling strategy for forecasting wildfire spread scenarios based on the assimilation of observed fire front location and on the sequential correction of model parameters or model state. This model relies on an estimation of the local rate of fire spread (ROS) as a function of environmental conditions based on Rothermel’s ROS formulation, in order to propagate the fire front with a level-set-based front-tracking simulator. In Part I, a data assimilation system based on an ensemble Kalman filter (EnKF) was implemented to provide a spatially-uniform correction of biomass fuel and wind parameters and thereby, produce an improved forecast of the wildfire behavior (addressing uncertainties in the input parameters of the ROS model only). In Part II, the objective of the EnKF algorithm is to sequentially update the two-dimensional coordinates of the markers along the discretized fire front, in order to provide a spatially-distributed correction of the fire front location and thereby, a more reliable initial condition for further model time-integration (addressing all sources of uncertainties in the ROS model). The resulting prototype data-driven wildfire spread simulator is first evaluated in a series of verification tests using synthetically-generated observations; tests include representative cases with spatially-varying biomass properties and temporally-varying wind conditions. In order to properly account for uncertainties during the EnKF update step and to accurately represent error correlations along the fireline, it is shown that members of the EnKF ensemble must be generated through variations in estimates of the fire initial location as well as through variations in the parameters of the ROS model. The performance of the prototype simulator based on state estimation or parameter estimation is then evaluated by comparison with data taken from a controlled grassland fire experiment. Results indicate that data-driven simulations are capable of correcting inaccurate predictions of the fire front location and of subsequently providing an optimized forecast of the wildfire behavior at future lead-times. The complementary benefits of both parameter estimation and state estimation approaches, in terms of
analysis and forecast performance, are also emphasized. In particular, it is found that the size of the assimilation window must be specified adequately with the persistence of the model initial condition and/or with the temporal and spatial variability of the environmental conditions in order to track sudden changes in wildfire behavior.

1 Introduction

Wildfires generally feature a front-like geometry and may be described at regional scales (i.e., at scales ranging from a few tens of meters up to several kilometers) as a thin flame zone that self-propagates normal to itself into unburnt vegetation. The local propagation speed is commonly referred to as the rate of spread (ROS) and is defined as the speed of the flame with respect to a fixed observer. Thus, the ROS can be regarded as the displacement rate of the flame surface separating the burning zone and the unburnt vegetation; it directly results from multi-scale multi-physics interactions between biomass pyrolysis, combustion and flow dynamics, radiation and convection heat transfer, as well as atmospheric dynamics and chemistry.

For early warning of potential wildfire danger, operational systems have been designed worldwide by national civil defense authorities to identify geographical areas that are subject to possibly extreme wildfire behavior. Fire danger is a generic term referring to the assessment of both constant and variable fire precursor factors affecting the ignition, spread, intensity and suppression capability of wildfires (Chandler et al., 1983). Current operational fire danger rating systems adopt a regional-scale viewpoint based on empirical ROS modeling approaches and integrate remote sensing data (i.e., meteorological, terrain topography and biomass fuel information) into a reduced set of macroscopic qualitative and/or numerical indices such as the Fire Weather Index (FWI) in Canada and in Europe (Van Wagner, 1987; Hirsch, 1996), the Forest Fire Danger Index (FFDI) in Australia (Noble et al., 1980; Dowdy et al., 2009) or the National Fire Danger Rating System (NFDRS) in the USA (Burgan, 1988). These fire danger rating systems mainly rely on meteorological information. Thus, the evaluation
of fire danger could be improved through a more detailed wildfire spread modeling and a more accurate forecast of the potential ROS, accounting for the spatial and temporal variability of environmental conditions.

A wide range of wildfire spread modeling approaches has emerged over the last decade to describe the physical processes at flame scale as well as the interactions between the fire and the atmosphere. On the one hand, physics-based computational fluid dynamics (CFD) modeling approaches provide detailed numerical simulations of the combustion-related processes that control the fire spread. Due to the high computational cost of flame-scale CFD and due to the lack of knowledge in environmental conditions, the use of CFD-based detailed modeling approaches such as FIRETEC (Linn et al., 2002), WFDS (Mell et al., 2007) or AVBP-PRISSMA-PYROWO (Rochoux, 2014) is currently restricted to research projects and is not compatible with operational applications. On the other hand, regional-scale fire spread models such as FARSITE (Finney, 1998), FOREFIRE (Filippi et al., 2009) or PHOENIX RapidFire (Chong et al., 2013) use a semi-empirical model that treats the ROS as a parametric function of biomass fuel properties, terrain topography and meteorological conditions; for instance, FARSITE uses a model due to Rothermel (1972); a detailed review of empirical and semi-empirical ROS models is provided in Sullivan (2009). One recent strategy to better account for time-varying weather conditions at regional scales consists in coupling a front-tracking simulator for surface fires with a meso-scale CFD atmospheric model for fire-induced atmospheric dynamics, see for instance WRF-Fire (Kochanski et al., 2013) or FOREFIRE-MESONH (Filippi et al., 2013). Still, many uncertainties remain due to simplifications in the description of the physics and to knowledge gaps in the description of environmental conditions and yet, errors in the properties of the biomass fuel or in the flame/wind interactions induce strong changes in the heat transfer from the flame to the vegetation and in the biomass fuel pyrolysis for instance. Thus, modeling errors as well as errors in the boundary conditions (e.g., biomass fuel properties, meteorological conditions and terrain topography) and initial conditions inevitably translate into errors in the simulated and forecast positions of the fire front.
For the wildfire spread simulation to be predictive and compatible with operational applications, the uncertainties in the ROS semi-empirical model must be quantified and reduced through a data assimilation (DA) methodology. The uncertainties inherent in wildfire spread modeling suggest the use of ensemble-based DA. Ensemble forecasts stochastically characterize the non-linear response of models to variations in the input data (Finney et al., 2011). Furthermore, DA integrates available fire sensor observations into a computer modeling tool to correct and optimize the model outputs and to thereby, produce improved forecast capabilities. Since sharp variations of the wind direction can lead to significant spatial deformations of the propagating fireline and since a canyon terrain can lead to eruptive fire behavior with strong acceleration in the steepest upslope directions, a spatially-distributed correction along the fireline is required to be able to track the time-evolving location of the fire front.

This study is an extension of our previous work presented in Rochoux et al. (2012, 2013a, b, 2014), in which a prototype data-driven wildfire simulator was developed. The initial prototype featured the following main components: a level-set-based fire propagation solver combined with a model description of the local ROS proposed by Rothermel (1972); an assumed series of airborne or possibly spaceborne observations of the fire front location; and a DA algorithm relying on parameter estimation. The DA prototype based on an ensemble Kalman filter (EnKF) was successfully evaluated in the context of a controlled grassland fire experiment (Paugam et al., 2013) when applied for estimating the input parameters used in the Rothermel-based ROS model (e.g., the fuel moisture content, the fuel particle surface-to-volume ratio, the wind direction and magnitude), see for instance Rochoux et al. (2014). It was found that data-driven simulations are capable of correcting inaccurate predictions of the fire front location and of providing an optimized forecast of the wildfire behavior; the quality of the forecast prediction capability being dependent on the temporal variability of the errors in the ROS model parameters. While these studies confirmed the potential of a DA strategy for improved wildfire spread predictions, the choice of a spatially-uniform parameter estimation approach is considered questionable. While well-suited for statis-
tically spatially-homogeneous problems (i.e., problems in which corrections to the parameters of the ROS model can be applied uniformly), this choice is no longer adapted to more general wildfire problems in which biomass fuel, topographical and possibly meteorological properties exhibit arbitrary spatial variations. In order to provide a spatial correction along the fireline, one could foresee the extension to the estimation of spatially-distributed biomass, topographical, and wind ROS model parameters. As this solution would be computationally prohibitive in the context of real-time forecast of wildfire behavior and inconsistent with the actual knowledge of environmental conditions, an alternative solution is proposed here.

The objective of the present study is to remove some of the main limitations in the initial design of our prototype data-driven wildfire simulator with an extension to the case of spatially-varying biomass fuel and meteorological conditions (the effect of terrain topology is outside the scope of the present study and a flat terrain is assumed). This extension is based on a change from a parameter estimation (PE) approach to a state estimation (SE) approach that consists in sequentially updating the model state using complete or incomplete observations of the fire front (see Fig. 1). This correction of the fire front location at the observation time provides a more reliable initial condition to the fire spread model, and allows to produce ensemble-based forecasts of the fire front location that are more consistent with the actual fire behavior. In contrast to the PE approach that only addresses uncertainties in the ROS model parameters, the SE approach accounts for multiple sources of uncertainties that are difficult to distinguish but that contribute to the uncertainties in the fire front location, i.e., uncertainties in the ROS model parameters but also in the ROS parameterization itself noted $f$ in Eq. (1):

$$\text{ROS} \equiv f(\text{wind conditions, biomass fuel properties}). \quad (1)$$

The uncertainties in the function $f$ are partly due to simplifications in the physics of propagating fires when building the relation between ROS and environmental conditions (e.g., heat transfer to biomass fuels, steady-state assumption, non-local effects).
This change to a SE approach was inspired in part by previous studies by Mandel and Beezley (2007) and Mandel et al. (2008, 2011) in which the control variable is the temperature state variable and is characterized by a bimodal probability density function (PDF) in the fire region (i.e., burning state or non-burning state). In order to satisfy the Gaussian assumption in the EnKF, the idea of morphing from image processing was introduced (Mandel and Beezley, 2007); however, this choice led to technical difficulties in the EnKF implementation. In the present study, the fire front is represented as a finite number of front markers, which are equally-spaced along the fireline that is projected onto a horizontal plane. Thus, the control vector includes the two-dimensional coordinates of each simulated front marker and is characterized by an approximate Gaussian PDF, which allows for a straightforward application of the EnKF. To obtain an efficient correction on the location of these front markers from a local observation, it is of primary importance to estimate reliable error correlations in the forecast error covariance matrix. Reliable means that these error correlations must be consistent with the actual spatial correlations of the errors in the fire spread model. For this purpose, an ensemble that accurately accounts for all sources of uncertainties must be generated during the EnKF prediction step.

The paper is organized as follows. The fire front observations and the wildfire spread model (called FIREFLY) are presented in Sect. 2. The EnKF algorithm is presented in Sect. 3. The performance of the resulting data-driven wildfire simulator is evaluated using first, academic tests in which observations are synthetically-generated in Sect. 3; the focus is on the impact of ensemble generation on the representation of error correlations and on the resulting shape of the corrected fireline. The simulator performance is evaluated using second, a validation test in which observations are taken from a controlled grassland fire experiment in Sect. 4; a comparative study of the SE approach with the PE approach is also provided to highlight the benefits and drawbacks of each approach, in terms of update and forecast performances.
2 Information on wildfires at regional-scales: observations and forward model

2.1 Observations of the fire front location

We assume in the present study that observations of the fire front location are available and that these observations can be made at different relevant times with a low measurement error. There is a growing body of literature on recent technological developments for geo-referenced wildfire front tracking (Riggan and Robert, 2009; Wooster et al., 2013), see Rochoux et al. (2014) for a comprehensive review.

In the following, the observation vector noted $y^o_t$ contains the two-dimensional coordinates $(x^o_i, y^o_i)$ of the front markers at time $t$, with $i$ varying between 1 and $N^o_{fr}$. The coordinates of the fire front markers are assumed to have independent Gaussian-like random errors $e^o_i$ with zero mean and with standard deviation (STD) noted $\sigma^o_i$. Two types of experiments are presented in this work: observation system simulation experiments (OSSE), in which observations are synthetically-generated using a reference (true) solution of the FIREFLY model that is modified by random observation errors $e^o_i$ (see Sect. 3); and a controlled grassland fire experiment, in which the observations are reconstructed from measured temperature maps and using a definition of the fire front as the 600 K iso-temperature contour (see Sect. 4).

2.2 The fire spread model (the forward model)

The fire spread model FIREFLY tracks the time-evolving location of the fire front using the following three components: (1) the Rothermel-based ROS model:

$$\Gamma \equiv \Gamma (\delta_v, M_v, M_{v, ext}, \Sigma_v, m''_v, \rho_p, \Delta h_c, u_w), \quad (2)$$

where the nomenclature for the input parameters are summarized in Table 1; (2) a level-set-based solver for the fire front propagation equation that computes the spatio-temporal variations of the progress variable $c$ introduced as a flame marker with $c = 0$ in the unburnt vegetation, $c = 1$ in the burnt vegetation, and where the flame front is...
identified as the isoline $c_{fr} = 0.5$; (3) an iso-contour algorithm for the reconstruction of the fire front that discretizes the flame contour $c_{fr} = 0.5$ with $N_{fr}$ markers, $(x_i, y_i)$ representing the two-dimensional coordinates of the $i$th marker. Note that the wind velocity magnitude $u_w$ is defined by the projection of the wind velocity vector (defined by its magnitude, $u_w^* \text{[m s}^{-1}]$, and direction angle, $\alpha_w^\circ$) along the (local) normal direction to the fire front, see Rochoux et al. (2014) for further details.

3 Data assimilation algorithm: specificities of the state estimation approach

3.1 Formulation of the ensemble Kalman filter

We present here the ensemble Kalman filter (EnKF) algorithm applied, in the context of SE, for one assimilation cycle between time $(t - 1)$ and time $t$.

3.1.1 Control space and observation space

The DA algorithm uses a discretization of both the simulated and observed fire fronts, called SFF and OFF, respectively. The discretization of SFF is a set of $N_{fr}$ markers; the control vector $x_t$, also called the state vector, contains the two-dimensional coordinates of the $N_{fr}$ front markers at time $t$:

\[ x_t = [(x_1, y_1), (x_2, y_2), \ldots, (x_{N_{fr}}, y_{N_{fr}})], \]

such that the size of $x_t$ is $n = 2N_{fr}$.

Similarly, the discretization of OFF is a set of $N_{fr}^o$ markers; the observation vector $y_t^o$ of size $2N_{fr}^o$ is defined as:

\[ y_t^o = \left[ (x_1^o, y_1^o), (x_2^o, y_2^o), \ldots, (x_{N_{fr}^o}, y_{N_{fr}^o}^o) \right]. \]
Observations of the fire front location are likely to be provided with a coarse resolution; in addition, observations may be incomplete and cover only a fraction of the fire front perimeter. Thus, we may expect $N_{fr}^o$ to be much lower than $N_{fr}$. In the following, we assume for simplicity that $N_{fr}^o = (N_{fr}/r)$, where $r$ is an integer taking values (much) larger than 1.

### 3.1.2 Observation operator

In order to pair a subset of $N_{fr}^o$ markers along SFF with the $N_{fr}^o$ markers along OFF, an observation operator $G_t$ is introduced. Within the framework of SE, $G_t$ is reduced to a selection operator $H_t$ taking 1 out of every $r = (N_{fr}/N_{fr}^o)$ markers along SFF at the observation time $t$. The model counterparts of the observation quantities, noted $y_t$ and of size $2N_{fr}^o$, read:

$$y_t = G_t(x_t) = H_t(x_t),$$

associating each marker of OFF with its closest neighbor along SFF.

### 3.1.3 Sequential estimation

The EnKF algorithm is sequentially applied; each assimilation cycle $[t - 1, t]$ consists in two successive steps for each member of the ensemble indexed by the exponent $k$ as illustrated in Fig. 2:

1. **a prediction step (forecast)**, in which the system is evolved from time $(t - 1)$ to time $t$ ($t$ being the next observation time) through an integration of FIREFLY to forecast the fire front location $x_t$ given some uncertainty ranges in the ROS model parameters and in the fire ignition location $(x_{ign}, y_{ign})$. This step leads to an ensemble of $N_e$ fire front positions at time $t$ designated as:

$$x_t^f = \left[ x_t^{f,(1)}, \ldots, x_t^{f,(k)}, \ldots, x_t^{f,(N_e)} \right],$$
used to stochastically describe the error covariance matrix $P^f_t \in \mathbb{R}^{2N_{fr} \times 2N_{fr}}$ that is expressed as:

$$P^f_t = \sum_{k=1}^{N_e} \frac{\left( x^f_t(k) - \bar{x}^f_t \right) \left( x^f_t(k) - \bar{x}^f_t \right)^T}{N_e - 1},$$

(4)

where the overline denotes the mean value over the ensemble. The structure of $P^f_t$ is as follows:

- The first diagonal block (of size $N_{fr} \times N_{fr}$) represents the error covariances of the marker $x$-coordinates (univariate error covariances): each diagonal term represents the error variance of one marker $x$-coordinate, while off-diagonal terms represent the covariances of the error in one marker $x$-coordinate with the errors in the other marker $x$-coordinates;

- The second diagonal block (of size $N_{fr} \times N_{fr}$) represents the error covariances of the $y$-coordinates (univariate error covariances): each diagonal term represents the error variance of one marker $y$-coordinate, while off-diagonal terms represent the covariances of the error in one marker $y$-coordinate with the errors in the other marker $y$-coordinates;

- Cross-diagonal blocks (of size $N_{fr} \times N_{fr}$) represent the (symmetric) multivariate error crossed-covariances between the $x$- and $y$-coordinates of the $N_{fr}$ simulated fire front markers (multi-variate error covariances).

2. an update step (analysis), in which new observations $y^o_t$ are considered at the analysis time $t$ and in which the $k$th ensemble member is updated consistently
with the observations, based on the following EnKF update equation:

\[
x^a_k(t) = x^f_k(t) + K^e_k \left( y^o_t + \xi^o_k - H_t \left( x^f_k(t) \right) \right),
\]

\[
K^e_k = P^f_t H^T_t \left( H_t P^f_t H^T_t + R \right)^{-1},
\]

where, as for the PE approach (see Rochoux et al., 2014), the innovation vector \(d^k_t\) is simply defined as the vector formed by the directed distances between the paired SFF-OFF markers for the \(k\)th ensemble member:

\[
d^k_t = y^o_t + \xi^o_k - y^f_t = y^o_t + \xi^o_k - H_t \left( x^f_k \right),
\]

with an additional noise \(\xi^o_k\) added to the observation vector to avoid ensemble collapse (Burgers et al., 1998). The EnKF update in Eqs. (5) and (6) provides corrected positions \(x^a_t\) for the \(N_{fr}\) simulated markers along the fireline at time \(t\), but there is no feedback on the ROS model parameters and/or on the fire initial location used to generate variability in the ensemble of forecasts. Note that in the context of SE, the selection operator \(H_t\) is straightforward to compute, meaning that a selection of lines and columns of \(P^f_t\) is sufficient to estimate the term \((H_t P^f_t H^T_t)\). Note also that in the present study as in Rochoux et al. (2014), observation errors are assumed uncorrelated, i.e., the observation error covariance matrix \(R\) is treated as a diagonal matrix, in which each diagonal term is the error variance \((\sigma^o)^2\) associated with the \(x\)- or \(y\)-coordinate of the \(N^o_{fr}\) markers along OFF.

### 3.2 Ensemble generation

The generation of the forecast ensemble is of primary importance in the performance of the EnKF forecast/update steps since it directly impacts how uncertainties are represented in the forecast error covariance matrix \(P^f_t\). To allow for reliable error covariance
modeling, all possible sources of uncertainties must be accounted for during this forecast step.

Different techniques can be considered to generate the ensemble of forecasts. If each fire front marker is perturbed separately (meaning that the error of one marker is uncorrelated to the errors of its neighbors along the fireline), the resulting fire front does not exhibit coherent features. This ensemble generation is therefore conflicting with the physical processes underlying wildfire spread. An alternative and more physically-consistent strategy is to generate an ensemble of simulated fire fronts by (1) randomly perturbing the input parameters of the Rothermel-based ROS model (e.g., wind magnitude \( u^*_w \) and direction angle \( \alpha^*_w \), fuel moisture content \( M_v \)) as well as the fire initial location \((x_{ign}, y_{ign})\), and by (2) integrating FIREFLY using the PALM-PARASOL functionality in Open-PALM\(^1\) (Fouilloux et al., 1999; Lagarde et al., 2001; Buis et al., 2006) for each set of parameters as in the PE approach, see Rochoux et al. (2014) for details on PALM-PARASOL. This leads to an ensemble of \( N_e \) fire front locations at time \( t \) designated as \( x^f_t \) and used to describe the forecast error covariance matrix \( P^f_t \).

A series of OSSE tests cases is presented here to highlight the key aspects of the proposed SE approach. The focus is on the impact of ensemble generation on the stochastic estimation of the spatial correlations of the marker location errors in \( P^f_t \) and the resulting correction of the fireline. Observation errors are assumed to be small with respect to the fireline perimeter (with the observation error STD taken equal to \( \sigma^0 = 1\) m); the SE-based EnKF performance is evaluated by its ability to track the time-evolving location of the observed fire front.

### 3.2.1 Isotropic case

An isotropic case corresponding to a 200 m x 200 m domain with uniform biomass fuel properties and no wind is considered first. The ROS \( \Gamma \) is constant, uniform and taken equal to \( 0.2\) m s\(^{-1}\). The true fire is ignited at \((x_{ign}, y_{ign}) = (100\) m, 100 m\) as a circular

\(^1\)http://www.cerfacs.fr/globc/PALM_WEB/
front with a radius of 5 m. Using a 1 m computational grid resolution along the \( x \)- and \( y \)-directions and a 0.5 s temporal resolution, FIREFLY is first integrated over the time window \([0; 200 \text{ s}]\) in order to produce at the analysis time (chosen to be \( t = 200 \text{ s} \)) the true location of the fire front. A forecast ensemble of \( N_e = 25 \) members is then produced based on spatial variations of the fire initial location \((x_{\text{ign}}, y_{\text{ign}})\) around a mean value \((97 \text{ m}, 103 \text{ m})\) and with an error STD along both \( x \)- and \( y \)-directions taken equal to \( \sigma_f = 10 \text{ m} \), see Fig. 3a. In this test, uncertainties in the forecast ensemble are only due to errors in the initial location of the fire front. The observed fire front is described by a stand-alone marker \((N_{\text{fr}}^0 = 1)\), while simulated fire fronts are discretized using \( N_{\text{fr}} = 100 \) markers (i.e., \( r = N_{\text{fr}}/N_{\text{fr}}^0 = 100 \)).

Figure 3a presents a comparison between the true and forecast fire front positions at time \( t = 200 \text{ s} \) (i.e., the observation time). This figure shows that due to uncertainties in the initial location of the fire front at initial time \((0 \text{ s})\), the predicted fire front locations are scattered over a large area at observation time. Since in this test, uncertainties in the distribution of the biomass fuel properties are not accounted for in the ensemble generation, the propagation of the fire front is isotropic (simulated fire fronts remain circular).

Figure 4 presents the spatial correlation of the error in the location of the marker indexed by \( m \) in Fig. 3a. Note that the error correlation associated with the marker \( m \) represents how the error in the location of this simulated marker is correlated to the errors in the location of its neighbors along the fireline (this, in order to characterize the fire front as a coherent feature given by the FIREFLY forward model). Figure 4 shows that univariate correlations are almost equal to one, while multi-variate correlations are nearly zero. As a consequence, errors in the position of the \( N_{\text{fr}} \) simulated front markers are highly correlated within the ensemble due to the isotropic propagation. As a result, the DA algorithm translates the information observed at one marker into a uniform correction along the fireline.

Figure 3b presents the comparison between the true and analysis fire front positions at the observation time \( t = 200 \text{ s} \). The analysis ensemble corresponds here to the up-
dated fire front locations that are produced through the EnKF update step. As expected, the analysis estimates of the fire front locations feature a much reduced scatter around the true location of the fire front, and the EnKF correction is isotropic due to high error correlations along the fireline (i.e., each analysis estimate within the ensemble is a circular front). With this uniform definition of the forecast ensemble, deforming the shape of the fire front through the ensemble-based analysis is therefore impossible.

### 3.2.2 Anisotropic case

To be able to stochastically represent more complex front shapes that are representative of all sources of uncertainties in the wildfire spread model and of their spatial variability, it is of primary importance to consider non-uniform environmental conditions when generating the ensemble of forecasts.

An anisotropic case of wildfire spread subject to spatially-varying vegetation properties and wind-aided propagation is now considered. An ensemble of $N_e = 20$ forecasts is produced over the time window $[0; 150 \text{s}]$, based on assumed uncertainties in the fire ignition location $(x_{\text{ign}}, y_{\text{ign}})$ at initial time (0 s) as well as in a subset of ROS model parameters, specifically in the fuel layer depth $\delta_v$, the fuel moisture content $M_v$, the fuel particle surface-to-volume ratio $\Sigma_v$ and the wind properties (magnitude $u^*_w$ and direction angle $\alpha^*_w$). In addition, the fuel depth $\delta_v$ is assumed to be spatially-varying, taking different values in the four quadrants of the square-shaped $700 \text{m} \times 700 \text{m}$ computational domain. Thus, uncertainties in the forecast ensemble are due to variations in 10 parameters, whose mean and STD values are presented in Table 2. Simulated fire fronts are discretized using $N_{\text{fr}} = 100$ markers; the impact of spatial error correlations on the EnKF-based analysis estimates is studied here for varying number of observed front markers $N_{\text{fr}}^0$.

Figure 5a presents a comparison between the true and forecast fire front positions at observation time $t = 150 \text{s}$. Due to uncertainties in the ROS model parameters and not only uncertainties in the fire initial location, the propagation is now anisotropic: the forecast fire fronts are characterized by stronger irregularities and more complex
shapes than results presented in Fig. 3a, due to the presence of wind and to the spatial variations in fuel depth. Figure 5b presents a similar comparison between the true and analysis fire front positions at time $t = 150s$; the observed fire front is described by $N_{fr} = 20$ uniformly-distributed markers (i.e., $r = 5$). As expected, the ensemble of analyses features a much reduced scatter (in terms of front shapes) around the true location of the fire front.

While Fig. 5a–b shows that the direct observation of the fire front location can overcome various uncertainties in the ROS model parameters, Fig. 5c illustrates that the spatial distribution of the observations along the fireline has a significant impact on the analysis. This figure considers a practically-relevant situation in which the observations are limited to a certain section of the fireline (i.e., the informed section, possibly due to the opacity of the thermal plume) and therefore, provide an incomplete picture. In this situation, while the EnKF algorithm produces an analysis that is close to the true state in the informed section, the benefits of DA are reduced in the non-informed sections. However, in spite of a reduced level of performance, the EnKF algorithm remains capable of a significant improved performance compared to a stand-alone forecast (in terms of front shapes and locations). This improved performance is due to the spatial correlation of the errors in the location of the simulated markers. Since the anisotropy of wildfire spread is now represented by a wide range of uncertainties in the forecast ensemble, the error in the location of one observed marker is correlated with the error in the location of its adjacent markers along the fireline as shown in Fig. 6 for the marker indexed by $m$ in Fig. 5a. Stated differently, the estimation problem translates the information coming from one observation marker into a local correction restricted to the closest neighbors only. The distance over which the observation marker affects the correction of the simulated front marker locations is referred to as correlation length-scale (Daley, 1991; Pannekoucke et al., 2013). This length-scale depends on the spatial variability of the errors in the ensemble generation and on the consistency of these statistics with the model errors statistics. For instance, Fig. 6 shows that the length-scale associated with the univariate error correlations of the marker $x$-coordinate typically takes values
on the order of 15 m on both sides of the considered simulated marker. This means that if this marker is assimilated, the correction of its location subsequently modifies the shape of the fire front over a distance of 15 m on both sides of this marker. This figure also shows that the errors in the marker y-coordinates are anti-correlated on both sides of the marker m: due to its particular location on the fireline (i.e., at the boundary between the flank and the back side of the fire), variations in the wind conditions induce significant changes in the fire front shape in the vicinity of marker m (markers can move from the flank to the back side of the fire or from the back side to the flank by modification of the wind direction for instance). As a consequence, when several observations are available, a non-uniform correction is obtained and the DA algorithm is able to change the shape of the analysis fronts and more easily match the observations.

Figure 7 examines the influence of the number of uniformly-distributed markers along the observed fire front $N_{fr}^O$ on the EnKF performance. This figure presents the root mean square (RMS) distance between the true and the forecast fire front positions as well as between the true and the analysis fire front positions as a function of $N_{fr}^O$. The figure shows that when $N_{fr}^O$ is large (see Fig. 5b with $N_{fr}^O = 20$), the EnKF algorithm successfully drives the analysis ensemble towards the true state; in contrast, when $N_{fr}^O$ is small, the EnKF algorithm has reduced effects and the analysis estimates remain close to the forecast estimates. Stated differently, the performance of the DA algorithm and its ability to capture the high-resolution features of wildfire spread depend strongly on the density of the observation network.

Furthermore, even though accounting for a wide range of uncertainties in the ROS parameters provides a wide range of possible fire front shapes, sampling errors can degrade the representation of the error statistics during the EnKF prediction step (if the number of members $N_e$ in the ensemble is not large enough) and thereby, the analysis solution. Figure 8 examines the impact of $N_e$ on the EnKF performance, in terms of the RMS distance between the true and forecast fire front positions as well as between the true and analysis fire front positions. These results show that for the present anisotropic case, $N_e = 20$ members are sufficient to obtain converged error
statistics; below this threshold value, sampling noise induces significant errors in the representation of error covariances.

### 3.3 Sequential model state estimation

#### 3.3.1 Principle and algorithm

To apply successive assimilations (i.e., perform regular EnKF updates), the ensemble of analysis fire fronts must be used as initial conditions and evolved to future times. However, to restart the FIREFLY front-tracking simulator, a two-dimensional progress variable field $c$ is required (see Sect. 2.2). The field $c^{(k)} \equiv c(x,y,t)^{(k)}$ associated with the $k$th ensemble member $x_t^{a,(k)}$ ($k$ varying between 1 and $N_e$) is therefore reconstructed and used as initial condition for the next assimilation cycle $[t, t + 1]$. This reconstruction is performed through a binarization process, i.e., $c^{(k)} = 0$ in the unburnt vegetation and $c^{(k)} = 1$ in the burnt area.

For each member $k$ in the ensemble, the reconstruction algorithm applied to each mesh node $(x_N, y_N)$ of the computational domain is as follows (see Fig. 9):

1. Pair the mesh node $(x_N, y_N)$ with the closest simulated front marker noted $(x_A, y_A)$ (computation of the minimal distance to the fire front);
2. Determine the closest neighbor of $(x_A, y_A)$ along the fireline noted $(x_B, y_B)$;
3. Compute the normal vector to the fireline (orthogonal to segment $[AB]$) passing through the node $(x_N, y_N)$, and determine their intersection noted $(x_O, y_O)$ (by linear algebra);
4. Calculate the two-dimensional inner product between the vector $\overrightarrow{ON}$ and the outward normal vector to the front at node $(x_A, y_A)$ noted $\overrightarrow{n_A}$: the sign of the inner product determines if the mesh node $(x_N, y_N)$ is inside or outside the burnt area:
if the inner product \( \langle \vec{ON} \cdot \vec{n} \rangle \) is positive, the mesh node \((x_N, y_N)\) is outside the burnt area;

- if \( \langle \vec{ON} \cdot \vec{n} \rangle \) is negative, the mesh node \((x_N, y_N)\) is within the burnt area;

- if \( \langle \vec{ON} \cdot \vec{n} \rangle \) takes a zero value, the mesh node \((x_N, y_N)\) is on the fireline.

5. Determine the corresponding value of the progress variable \( c \) at mesh node \((x_N, y_N)\); a smoothing hyperbolic tangent function is applied when the mesh node \((x_N, y_N)\) is near the fire front.

The reconstructed progress variable field \( c^{(k)} \) can be used to integrate FIREFLY over the next forecast period \([t; t + 1]\). A schematic of the sequential SE approach for the \( k \)th ensemble member is presented in Fig. 10; the reconstruction procedure must be performed for all members \((N_e)\) in the ensemble.

### 3.3.2 Sequential assimilation for the anisotropic case

Multiple assimilation cycles for the OSSE anisotropic case (presented in Sect. 3.2.2) are now considered based on the previously-described multi-cycle algorithm. The quality of the forecast (measured by its ability to track the location of the true fire front) is examined over the time window \([0; 600 \text{ s}]\) with observations taken at 150 s time intervals, at times \( t_1 = 150 \text{ s}, t_2 = 300 \text{ s}, t_3 = 450 \text{ s} \) and \( t_4 = 600 \text{ s} \). Thus, the EnKF update is performed over 4 successive assimilation cycles; each assimilation cycle \([t_{n-1}; t_n]\) includes a prediction step that integrates FIREFLY from \( t_{n-1} \) until \( t_n \) and an update step that corrects the location of the simulated fire front at time \( t_n \) \((n = 1, \ldots, 4)\). The propagation of the true fire front is simulated for time-varying wind conditions (but with assumed constant wind velocity and direction over each assimilation cycle) presented in Table 3, while the forecast ensemble is simulated for constant wind conditions using the ROS model parameters presented in Table 2. Note that the perturbation of the fire ignition location \((x_{\text{ign}}, y_{\text{ign}})\) is only introduced during the first assimilation cycle (as
a means to account for uncertainties in the fire initial location before remote sensing detection).

The quality of the forecast is expected to deteriorate at increasing lead-times (i.e., when the time delay between the actual forecast time and the previous analysis time increases) for two reasons. First, because the impact of the fire front correction applied at a given analysis time decreases with time (i.e., when the forecast lead-time increases). Second, because the present implementation of the SE-based EnKF does not provide any correction for ROS modeling errors or for ROS model parameters (including the incorrect assumption of a constant wind); such correction may be addressed through a PE approach, see Rochoux et al. (2014).

Figure 11 presents the successive comparison between the mean free run (i.e., stand-alone FIREFLY simulation without DA in green dashed-dotted line), the mean fire front location related to the forecast estimates (blue dashed line) and to the analysis estimates (red solid line) as well as the observations (black crosses) at times $t_1$, $t_2$, $t_3$ and $t_4$. It is found that the free run simulation does not accurately estimate the rate and direction of the fire spread due to imperfect knowledge in the ROS model parameters and in the fire initial location. As for the forecast, it provides a more accurate prediction of the fire front location and a more physically-consistent front shape than the free run at each assimilation time. Still, the analysis estimates exhibit a much reduced scatter due to the EnKF update and the information gain obtained by the observations: the distance between the analysis fire front and the observations is largely reduced at the analysis time and the shape of the analysis fire front is much more consistent with the observed fire front.

The performance of the EnKF update is confirmed by the error statistics presented in Fig. 12a, which presents the time-evolution of the deviations of FIREFLY model predictions from observations. It is shown that the benefits from the EnKF update decrease as the forecast lead-time increases. For instance, the RMS distance to the true front is significantly reduced by the EnKF update, from 30 m for the free run (FR) to less than 1 m for the analysis (A1) during the first assimilation cycle [0; 150 s]. Starting again
from the analysis estimates at time 150 s, FIREFLY simulates the forecast fire evolution in time but, without additional observations, the distance between the true state and the forecast (F1) significantly increases, to 20 m at time \( t_2 = 300 \text{s} \) up to approximately 80 m at time \( t_5 = 750 \text{s} \). By repeating the EnKF update at 150 s time intervals, the distance between the true state and the simulated fire front remains below 10 m. EnKF-based data-driven simulations bring valuable information on the wildfire spread behavior, even when DA is not applied systematically. This is illustrated in Fig. 12b, which presents a comparison of different fire front forecasts at time \( t_4 = 600 \text{s} \) using FIREFLY, with or without DA. The closer the assimilation time to \( t_4 = 600 \text{s} \), the better the forecast prediction: consistently, the free run (FR) provides the less accurate prediction with a mean distance to the observations approximately equal to 70 m; in contrast, this mean distance is reduced by a factor of nearly 2 for (F2) and by a factor of 70 for (A4).

In summary, these results on OSSE test cases show that in a SE approach, EnKF updates provide valuable information and lead to accurate forecasts on wildfire behavior at short lead-times, the definition of *short* being dependent on the persistence of the model initial condition. Accordingly, the EnKF update must be performed at regular time intervals to efficiently track the actual wildfire propagation.

### 4 Application to a controlled grassland fire experiment

The EnKF-FIREFLY data-driven simulator is now evaluated in a validation study corresponding to a controlled grassland fire experiment (a case in which the true control vector is not known and may not exist if the model is not sufficiently representative). The experimental configuration corresponds to a small \( 4 \text{m} \times 4 \text{m} \), flat and and horizontal, open-field grassland lot burning under moderate wind conditions in which the ROS takes values on the order of \( 1 \text{ cm s}^{-1} \) (with a maximum value reaching \( 5 \text{ cm s}^{-1} \) in the wind direction).
The properties of the grass are (approximately) known: $\delta_v = 8\, \text{cm}$ (field measurement), $M_v = 22\%$ (field measurement) and $\Sigma_v = 11\,500\, \text{m}^{-1}$ (Rothermel’s fuel database for short grass); the wind conditions are also approximately known: the magnitude and direction angle of the wind are constant and equal to $u_w^* = 1\, \text{m} \, \text{s}^{-1}$ and $\alpha_w^* = 307^\circ$. The fire spread is recorded using a thermal-infrared camera; the thermal maps are post-processed (the fire front is defined at the 600 K iso-temperature contour) and thereby, provide full fire contours. The study considers 4 successive, 14 s-long, assimilation cycles with initialization at time $t_0 = 50\, \text{s}$ and successive updates at times $t_1 = 64\, \text{s}$, $t_2 = 78\, \text{s}$, $t_3 = 92\, \text{s}$ and $t_4 = 106\, \text{s}$. Based on the spatial resolution of the camera, the estimated STD of the measurement error is $\sigma^o = 0.05\, \text{m}$.

### 4.1 Performance of state estimation

To generate the ensemble of forecasts, the mean values of the ROS model parameters are the (approximately) known short grass properties and wind conditions. Spatially-uniform distributions for $M_v, \Sigma_v, u_w^*$ and $\alpha_w^*$ are assumed; in addition, the grassland lot is divided into 4 equally-sized rectangular-shaped sections and the fuel depth $\delta_v$ is treated as different in each zone, 6, 8, 10 and 12 cm from east to west (noted $\delta_{v,i}$, with $i = 1, \ldots, 4$, respectively). The uncertainty in the initial position of the fire at time $t_0$ (along the $x$-direction) is also accounted for; the mean is the observed fire front location taken from the experiment. The ensemble contains $N_e = 50$ members corresponding to 9 different choices of the ROS model parameters and of the fire initial location, i.e., $[\delta_{v,1}, \delta_{v,2}, \delta_{v,3}, \delta_{v,4}, M_v, \Sigma_v, u_w^*, \alpha_w^*, x_{\text{ign}}]$. The corresponding stochastic perturbations are characterized by relatively large levels as presented in Table 4. In these simulations, $N_{fr} = 100$ markers are used to represent simulated fire fronts and $N_{o fr} = 50$ markers are considered for observations (i.e., $r = 2$).

Figure 13a presents a comparison between the mean (ensemble-average) forecast estimate of the fire front location (as predicted by FIREFLY starting from time $t_0 = 50\, \text{s}$), the observations and the mean analysis estimate at time $t_1 = 64\, \text{s}$. It is seen that the mean free forecast (without DA) significantly underestimates the observed ROS of the
fire. In contrast, the predictions that are made after an EnKF update performed at time $t_1$ successfully reduce the distance between predictions and observations. In particular, the mean analysis estimate features a topology that is very consistent with the observed fire front, a result that requires an accurate and non-uniform correction of the fire front location. Figure 14 presents the univariate error correlations associated with the front marker indexed by $m$ in Fig. 13a for the forecast (blue dashed line) and analysis (red solid line) estimates. The analysis estimates feature a much reduced length-scale compared to the forecast estimates to allow for a spatially-distributed correction during the EnKF update.

Similar comparisons to Fig. 13a are presented in Fig. 13b–d at times $t_2$, $t_3$ and $t_4$, respectively. Note that the mean forecast estimate (blue dashed line) is initialized at the previous observation time by the analysis produced by the DA cycle, while the free forecast (green dashed-dotted line) is initialized at time $t_0$ and does not use any analysis. It is found that the mean distance to the observations is reduced by a factor of at least 5 over all assimilation cycles, by performing EnKF updates at 14 s time intervals. Thus, the agreement between predicted and observed front positions is remarkable and significantly better than the level of agreement that would be obtained in the absence of DA. However, in spite of the quality of the correction provided by the SE-based EnKF algorithm, the performance of the forecast remains limited: for instance, Fig. 13b shows that the mean forecast estimate (initialized by the analysis at time $t_1$ and integrated until time $t_2$), while still significantly more accurate than the free forecast (initialized by the initial conditions at time $t_0$ and integrated until time $t_2$), is not in good agreement with the observation (relatively to the mean analysis estimate). These results suggest that in the present configuration, the persistence of the initial condition is limited to short lead-times (i.e., shorter than 14 s) and the performance of the SE-based EnKF forecast could be improved by more frequent assimilation.
4.2 Comparison to parameter estimation

A PE approach based on the standard EnKF approach and validated in Rochoux et al. (2014) is applied to the same controlled grassland fire experiment. While the SE approach leads to a direct adjustment of the fire front location, the PE approach works by an adjustment of the ROS model parameters. In the PE approach, 4 parameters are used as control variables: the fuel moisture content $M_v$ and particle surface-to-volume ratio $\Sigma_v$ as well as the wind magnitude $u_w^*$ and direction angle $\alpha_w^*$, i.e.,

$$x = [M_v, \Sigma_v, u_w^*, \alpha_w^*]$$

with $n = 4$ the size of the control vector. These parameters are perturbed around mean values and with prescribed uncertainties (according to assumed levels of uncertainty presented in Table 4), but remain spatially-uniform for each ensemble member. The EnKF ensemble contains $N_e = 1000$ members, meaning that during each assimilation cycle, FIREFLY produces 1000 fire front trajectories associated with each set of control parameters. Note that the size of the ensemble is drastically increased compared to the SE approach, in order to retrieve accurate error statistics of the control parameters and to avoid the equifinality problem (i.e., a problem in which multiple sets of parameters provide the same simulated fire front), see Rochoux (2014) for further explanations on this equifinality problem.

Figure 15a compares the mean forecast estimate obtained at time $t_4 = 106s$, using the PE ensemble (blue dashed line) and the SE ensemble (blue solid line) as predicted by FIREFLY integration and with an EnKF update performed at successive observation times $t_1 = 64s$, $t_2 = 78s$ and $t_3 = 92s$. Figure 15b compares the mean analysis estimate obtained at the same time $t_4 = 106s$ using a PE-based EnKF update (red dashed line) and a SE-based EnKF update (red solid line). In both figures, predictions are compared to observations (black crosses). It was found that for both PE and SE approaches the analysis estimates provide more accurate fire front locations than the forecast estimates. This point is also illustrated in Fig. 16, which presents the RMS
distance between the observations and the mean fire front location produced by the forecast estimates (see Fig. 16a) and analysis estimates (see Fig. 16b) over the 4 successive assimilation cycles: the mean distance to the observations is reduced by a factor 2 in the PE approach and by a factor of at least 5 in the SE approach. Additionally, the performance of the SE-based analysis is significantly better than that of the PE-based analysis (relying on spatially-uniform distributions of ROS model parameters): the mean distance between observed and SE-based simulated fire fronts remains below 0.1 m for all assimilation cycles and the scatter of the SE-based analysis ensemble is significantly less than that obtained through PE.

In spite of the overall quality of the correction provided by both EnKF estimation approaches, the accuracy of the forecast, while still significantly better than that obtained in the free run simulations, rapidly decreases over time. Some of the benefits of an analysis are indeed lost in the forecast at time $t_4$, due to the limited persistence of the initial condition and/or due to the temporal variability of the errors in the environmental conditions. The improved forecast performance of the PE approach is illustrated in Fig. 16a; the adjustment of the ROS model parameters allows for a correction of inaccuracies in initial guesses as well as an adaptation to time-dependent properties; the statistical properties of the EnKF ensemble are dynamically-evolving. In contrast, the statistical properties of the SE-based EnKF ensemble are constant. Figure 17 illustrates this point. In particular, this figure shows that the prior values of the moisture content $M_v$ and the fuel particle surface-to-volume ratio $\Sigma_v$ are not adequate to effectively track the actual fire propagation (in comparison, the estimated wind conditions in the PE approach are contained in the support of the forecast PDFs). These values are corrected in the PE approach: the mean (ensemble-average) value of $M_v$ is decreased from 22% to approximately 10%; the mean (ensemble-average) value of $\Sigma_v$ is increased from $11500 \text{m}^{-1}$ to approximately $19000 \text{m}^{-1}$. These values are not corrected in the SE approach and therefore, induce a significant bias in wildfire spread simulations.
In this validation study, the assimilation needs to be renewed by frequent observations with an assimilation frequency (i.e., frequency at which the EnKF update is renewed by observations) that is high enough to track the temporal variability of the errors in the ROS model parameters. Updating environmental conditions (the biomass moisture content in particular) explains the bulk of the improved forecasts at future lead-times.

5 Conclusions

This study is the second part in a series of two articles, in which we propose and explore a new paradigm for producing optimized forecasts of the wildfire behavior based on observations of the fire front location. The prototype data-driven simulator combines a regional-scale wildfire spread model, FIREFLY, with a data assimilation algorithm based on an ensemble Kalman filter (EnKF); it features a choice between a parameter estimation (PE) approach (in which the estimation targets are the parameters of the rate of spread – ROS – model) and a state estimation (SE) approach (in which the estimation targets are the time-evolving spatial coordinates of the fire front). The cornerstone of this prototype data-driven simulator is to (1) explicitly account for the effects of both measurement and modeling errors and thereby, overcome some of the current limitations of regional-scale wildfire modeling; (2) account for non-linearities in the wildfire behavior and for temporal variability of the errors in the environmental conditions; and (3) forecast reliable wildfire spread scenarios at limited computational cost, consistently with an operational framework for real-time monitoring of wildfire behavior. The present study assumes that airborne and/or spaceborne observations of the fire front location are available at frequent times but possibly provide an inaccurate and incomplete description of the fire front due to the opacity of the fire-induced thermal plume or due to a limited monitoring. The performance of both PE and SE approaches was evaluated on synthetic cases of wildfire spread that are representative of real wildfire conditions as well as on a controlled grassland burning experiment.
Even though certain input parameters of the ROS model can be assumed constant over the fire duration (in particular, intrinsic properties of biomass fuels), other parameters exhibit a dynamic behavior due to the presence of the propagating wildfire. Wind conditions change over time, partly due to the fire/atmosphere interactions and their feedback on local atmospheric conditions (in terms of wind, air temperature and humidity). In addition, the moisture content of biomass fuel also varies, in particular that of dead biomass, which can be considered in thermal equilibrium with the atmosphere. A dynamic estimation of time-varying environmental parameters is therefore required to produce accurate simulations of the wildfire behavior. For this purpose, a PE approach was presented in Part I of this series of two articles, see Rochoux et al. (2014). This PE approach was restricted to spatially-uniform parameters. Accounting for the detailed spatial variations of environmental conditions would indeed significantly increase the computational cost, with no means of assessing the consistency of the EnKF update to in-situ measurements (since those are usually provided with a coarse spatial resolution).

Part II of this series of two articles demonstrated that to obtain a local correction of the fire front location, the generation of the stochastic ensemble must represent the anisotropy in wildfire spread. This anisotropy was implicitly introduced in the SE approach by selecting spatially-distributed biomass fuel properties and distinct wind conditions between the different members in the ensemble. The SE approach was shown to be successful at retrieving at low computational cost (i.e., with a relatively small number of members in the ensemble) the actual shape of the fire front in cases with strong anisotropic propagation conditions. This approach was also found relevant for observations made with significant errors and/or cases in which the observations are incomplete (e.g., when only a section of the fireline is informed), in order to reconstruct a complete, reliable initial condition for FIREFLY restart.

Furthermore, it was demonstrated that for the present grassland controlled fire experiment, the PE approach provides a more reliable forecast capability of the wildfire behavior than the SE approach (due to improved knowledge in environmental condi-
tions that are usually poorly assessed) at short lead-times (i.e., at 14 s time intervals). This ranking between PE and SE approaches remains, however, problem-dependent and may not hold in cases where the assimilation cycle is longer (due to a lower observation frequency) and where the values of the environmental conditions strongly fluctuate during an assimilation cycle. The duration of the assimilation cycle (i.e., a user-defined variable) is therefore of primary importance in the success of the proposed data assimilation approaches; it must be specified consistently with the persistence of the initial condition and/or with the temporal and spatial variability of the errors in the environmental conditions.

This series of two articles emphasizes the potential of data assimilation to dramatically increase wildfire simulation accuracy. While wildfire spread forecast capabilities are still at an early stage of development, it is envisioned that they will be similar to current weather forecast capabilities and that they will provide real-time fire forecasts using thermal-infrared imaging including a description of both wildfire dynamics and plume emissions such as FOREFIRE-MESONH (Filippi et al., 2009, 2013). Thus, future plans include the development of a dual SE/PE approach that could partly overcome the forecast limitations illustrated in the present and past studies (the PE approach could be extended to the case of weak spatial variations of the ROS model parameters). Future plans also include the integration of the data assimilation algorithm into a meso-scale atmospheric model based on computational fluid dynamics. The issues related to data assimilation for coupled physical systems were not addressed in this work. We propose here a data assimilation strategy that is suitable for surface wildfire spread, but that would need further developments to deal with a land-surface–atmosphere modeling system.

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\textsuperscript{2}http://anridea.univ-corse.fr
perimental data, Florent Duchaine and Thierry Morel (CERFACS) for support on Open-PALM, as well as Jonathan Beezley for helpful discussions on state estimation problems for wildfire spread.

References


Table 1. Input parameters of the Rothermel-based ROS sub-model.

<table>
<thead>
<tr>
<th>Name</th>
<th>Symbol</th>
<th>Unit</th>
</tr>
</thead>
<tbody>
<tr>
<td>Fuel depth (vertical thickness of the vegetation layer)</td>
<td>$\delta_v$</td>
<td>m</td>
</tr>
<tr>
<td>Fuel moisture (mass of water divided by mass of dry vegetation)</td>
<td>$M_v$</td>
<td>%</td>
</tr>
<tr>
<td>Fuel moisture at extinction</td>
<td>$M_{v,ext}$</td>
<td>%</td>
</tr>
<tr>
<td>Fuel particle surface-to-volume ratio</td>
<td>$\Sigma_v$</td>
<td>m$^{-1}$</td>
</tr>
<tr>
<td>Fuel loading</td>
<td>$m''_v$</td>
<td>kg m$^{-2}$</td>
</tr>
<tr>
<td>Fuel particle mass density</td>
<td>$\rho_p$</td>
<td>kg m$^{-3}$</td>
</tr>
<tr>
<td>Fuel heat of combustion</td>
<td>$\Delta h_c$</td>
<td>J kg$^{-1}$</td>
</tr>
<tr>
<td>Wind velocity magnitude at mid-flame height (projected onto horizontal plane)</td>
<td>$u_w$</td>
<td>m s$^{-1}$</td>
</tr>
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</table>
### Table 2. Properties of the ensemble forecast in the OSSE anisotropic case.

<table>
<thead>
<tr>
<th>Random variable</th>
<th>True value</th>
<th>Ensemble mean</th>
<th>Ensemble STD</th>
</tr>
</thead>
<tbody>
<tr>
<td>$(\delta_{v,1}, \delta_{v,2}, \delta_{v,3}, \delta_{v,4})$ [m]</td>
<td>(0.25, 1.25, 0.75, 1.75)</td>
<td>(0.25, 1.25, 0.75, 1.75)</td>
<td>(0.10, 0.10, 0.10, 0.10)</td>
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<tr>
<td>$M_v$ [%]</td>
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<td>20</td>
<td>10</td>
</tr>
<tr>
<td>$\Sigma_v$ [m$^{-1}$]</td>
<td>11 500</td>
<td>11 500</td>
<td>4000</td>
</tr>
<tr>
<td>$(u_{w}^<em>, \alpha_{w}^</em>)$ [m s$^{-1}$, °]</td>
<td>(1.0, 315)</td>
<td>(0.75, 315)</td>
<td>(0.15, 45)</td>
</tr>
<tr>
<td>$x_{ign}$ [m]</td>
<td>350</td>
<td>350</td>
<td>20</td>
</tr>
<tr>
<td>$y_{ign}$ [m]</td>
<td>350</td>
<td>350</td>
<td>20</td>
</tr>
</tbody>
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Table 3. Time-varying true wind conditions (in terms of magnitude $u^*_w$ and direction angle $\alpha^*_w$) in the OSSE anisotropic case with multiple assimilation cycles, from 0 to 600 s, at 150 s time intervals.

<table>
<thead>
<tr>
<th>Assimilation cycle</th>
<th>$u^*_w$ [m s$^{-1}$]</th>
<th>$\alpha^*_w$ [$^\circ$]</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>1.0</td>
<td>315</td>
</tr>
<tr>
<td>2</td>
<td>0.75</td>
<td>290</td>
</tr>
<tr>
<td>3</td>
<td>0.83</td>
<td>257</td>
</tr>
<tr>
<td>4</td>
<td>1.20</td>
<td>232</td>
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</table>
Table 4. Statistical properties of the forecast ensemble in the controlled grassland fire experiment.

<table>
<thead>
<tr>
<th>Random variable</th>
<th>Ensemble mean</th>
<th>Ensemble STD</th>
</tr>
</thead>
<tbody>
<tr>
<td>$M_v$ [%]</td>
<td>22</td>
<td>6</td>
</tr>
<tr>
<td>$\Sigma_v$ [m$^{-1}$]</td>
<td>11 500</td>
<td>4000</td>
</tr>
<tr>
<td>$(u^<em>_w, \alpha^</em>_w)$ [m s$^{-1}$, °]</td>
<td>(1, 307)</td>
<td>(0.4, 45)</td>
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<tr>
<td>$\delta_{v,1}$ [m]</td>
<td>0.06</td>
<td>0.04</td>
</tr>
<tr>
<td>$\delta_{v,2}$ [m]</td>
<td>0.08</td>
<td>0.04</td>
</tr>
<tr>
<td>$\delta_{v,3}$ [m]</td>
<td>0.10</td>
<td>0.04</td>
</tr>
<tr>
<td>$\delta_{v,4}$ [m]</td>
<td>0.12</td>
<td>0.04</td>
</tr>
<tr>
<td>$x_{ign}$ [m]</td>
<td>2</td>
<td>0.65</td>
</tr>
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</table>
Figure 1. Data assimilation flowchart for PE and SE approaches over one assimilation cycle (control variables are colored in blue).
Figure 2. Flowchart of the EnKF algorithm during the $[t-1, t]$ assimilation cycle for a SE approach.
(a) Forecast estimates of the fire front location.

(b) Analysis estimates of the fire front location.

**Figure 3.** Spatially-uniform OSSE test with constant ROS but uncertain ignition location ($x_{\text{ign}}, y_{\text{ign}}$); one assimilation cycle; all figures correspond to time $t = 200$ s. **(a)** Comparison between true (black solid line) and forecast (blue dashed lines) fire front positions; the cross symbol is the only observation available. **(b)** Similar comparison between true (black solid line) and analysis (red dashed lines) fire front positions.
Figure 4. Error correlation functions along the fire front associated to one marker location (the vertical bar represents its reference location indexed by m in Fig. 3a) for the OSSE isotropic propagation test with constant ROS but uncertain ignition location ($x_{ign}$, $y_{ign}$). (a) Univariate correlations: the dashed (solid) line indicates the error correlation of the reference marker $x$-coordinate ($y$-coordinate) with respect to the errors in the $x$-coordinates ($y$-coordinates) of the other markers along the fireline. (b) Multi-variate correlations: the dashed (solid) line indicates the error correlation of the reference marker $x$-coordinate ($y$-coordinate) with respect to the errors in the $y$-coordinates ($x$-coordinates) of the other markers along the fireline.
(a) Forecast estimates of the fire front location.
(b) Analysis estimates of the fire front location with $N_{fr}^o = 20$.
(c) Analysis estimates of the fire front location with $N_{fr}^o = 12$ for an incomplete set of observations.

Figure 5. Spatially-varying OSSE test with uncertain ROS model parameters and with uncertain ignition location ($x_{ign}, y_{ign}$); one assimilation cycle; all figures correspond to time $t = 150$ s. (a) Comparison between true (black solid line) and forecast (blue dashed lines) fire front positions. (b, c) Comparison between true (black solid line) and analysis (red dashed lines) fire front positions. Black cross symbols are the $N_{fr}^o$ observed front markers.
(a) Univariate error correlations along the fireline.

(b) Multi-variate error correlations along the fireline.

**Figure 6.** See caption of Fig. 4. OSSE anisotropic case with uncertain ROS model parameters and with uncertain ignition location \((x_{\text{ign}}, y_{\text{ign}})\); the vertical bar represents its reference location indexed by \(m\) in Fig. 5a.
Figure 7. Average distance between the true and forecast fire front positions (blue circles) and between the true and analysis fire front positions (red squares) at time 150 s as a function of the number of observed markers $N_{fr}^o$; OSSE anisotropic case. Error bars indicate the associated STD.
Figure 8. Average distance between the true and forecast fire front positions (blue circles) and between the true and analysis fire front positions (red squares) at time 150 s as a function of the number of members $N_e$ in the ensemble; OSSE anisotropic case. Error bars indicate the associated STD.
Figure 9. Schematic of the reconstruction algorithm for the FIREFLY initial condition. In this example, the mesh node \((x_N, y_N)\) is outside the burnt area, the inner product \((\overrightarrow{ON} \cdot \overrightarrow{n_A})\) is positive.
Figure 10. Schematic of the sequential SE approach for each ensemble member \( k = 1, \cdots, N_e \), with a focus on the reconstruction of the two-dimensional progress variable \( c^{(k)} \) posterior to analysis over the time window \([t - 1, t]\) and prior to forecast over the time window \([t, t + 1]\).
Figure 11. OSSE anisotropic case over multiple assimilation cycles (from 0 to 600 s): comparison between the free run (green dashed-dotted line), mean forecast estimate (blue dashed line), observations (black crosses) and mean analysis estimate (red solid line) at 150 s time intervals, at times (a) $t_1 = 150$ s (the gray circle corresponds to the true initial location); (b) $t_2 = 300$ s; (c) $t_3 = 450$ s; and (d) $t_4 = 600$ s.
Figure 12. OSSE anisotropic case over multiple assimilation cycles (from 0 to 600 s). (a) Average distance between the true and forecast/analysis fire front location as a function of the assimilation cycle index: green circles correspond to the free run (FR); triangles, crosses, diamonds and stars correspond to a forecast with an EnKF update at $t_1 = 150 \text{s}$ (F1), $t_2 = 300 \text{s}$ (F2), $t_3 = 450 \text{s}$ (F3) and $t_4 = 600 \text{s}$ (F4), respectively; square symbols correspond to an EnKF update performed at times $t_1$ (A1), $t_2$ (A2), $t_3$ (A3), $t_4$ (A4) and $t_5$ (A5). (b) Comparison at time $t_4 = 600 \text{s}$ between the free run (FR) in green dashed-dotted line, the mean forecast estimate (F1, F2, F3) in dashed lines, observations in black crosses, and the mean analysis estimate (A4) in red solid line.
Figure 13. Controlled grassland fire experiment with multiple assimilation cycles from $t_0 = 50\,\text{s}$ to $t_4 = 106\,\text{s}$. Black crosses correspond to observations, the gray solid line corresponds to the initial condition of the assimilation cycle, the green dashed-dotted line corresponds to the free run (without DA), the blue dashed line corresponds to the mean forecast estimate (without DA for the first observation time or with DA at the previous observation time), and the red solid line corresponds to the mean analysis estimate (with a DA update at the current observation time). Times (a) $t_1 = 64\,\text{s}$; (b) $t_2 = 78\,\text{s}$; (c) $t_3 = 92\,\text{s}$; and (d) $t_4 = 106\,\text{s}$. 
Figure 14. Univariate error correlation functions along the fireline associated to one marker location (the vertical bar represents its reference location indexed by m in Fig. 13a); controlled grassland fire experiment over the assimilation cycle [50; 64s]. (a) Error correlation of the reference marker x-coordinate with respect to the errors in the x-coordinates of the other markers along the fireline: the blue dashed (red solid) line corresponds to the mean forecast (analysis) estimate. (b) Error correlation of the reference marker y-coordinate with respect to the errors in the y-coordinates of the other markers along the fireline: the blue dashed (red solid) line corresponds to the mean forecast (analysis) estimate.
Figure 15. Comparison between simulated (lines) and observed (symbols) front positions at time $t_4 = 106$ s in the controlled grassland burning experiment. The simulated front position is the mean position calculated as the average of the EnKF ensemble; dashed lines (solid lines) correspond to the PE-based (SE-based) simulations. (a) Forecast (with EnKF update at $t_3 = 92$ s). (b) Analysis (with EnKF update at $t_4 = 106$ s).
Figure 16. Evolution over 4 successive assimilation cycles of the average distance between the observations and the mean fire front location produced by the (a) forecast estimates and the (b) analysis estimates at times $t_1 = 64\,\text{s}$, $t_2 = 78\,\text{s}$, $t_3 = 92\,\text{s}$, and $t_4 = 106\,\text{s}$. Dashed (solid) lines represent PE (SE) results; vertical bars indicate the STD in the forecast and analysis estimates.
Figure 17. Evolution of ROS model parameters at 4 successive analysis times; controlled grassland fire experiment. PE-based ensemble: blue circle symbols (connected by a dashed line) indicate mean values in the forecast ensemble; red square symbols (connected by a solid line) indicate mean values in the analysis ensemble; vertical bars indicate the corresponding STD. SE-based ensemble: black horizontal solid lines indicate mean values; black horizontal dashed lines indicate STD. (a) Biomass fuel moisture content $M_v$ [%]. (b) Biomass fuel particle surface-to-volume ratio $\Sigma_v$ [m$^{-1}$]. (c) Wind magnitude $u'_w$ [m s$^{-1}$]. (d) Wind direction $\alpha'_w$ [°].